

1 Asymmetric competition causes multimodal size
2 distributions in spatially-structured populations

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16 **Abstract**

17 Plant sizes within populations often exhibit multimodal distributions, even when
18 all individuals are the same age and have experienced identical conditions. To
19 establish the causes of this we created an individual-based model simulating the
20 growth of trees in a spatially-explicit framework, which was parameterised using
21 data from a long-term study of forest stands in New Zealand. First we demonstrate
22 that asymmetric resource competition is a necessary condition for the formation of
23 multimodal size distributions within cohorts. In contrast, the legacy of small-scale
24 clustering during recruitment is transient and quickly overwhelmed by density-
25 dependent mortality. Complex multi-layered size distributions are generated when
26 established individuals are restricted in the spatial domain within which they can
27 capture resources. The number of modes reveals the effective number of direct
28 competitors, while the separation and spread of modes are influenced by distances
29 among established individuals. Asymmetric competition within local neighbour-
30 hoods can therefore generate a range of complex size distributions within even-aged
31 cohorts.

32 **Keywords**

33 bimodality; individual-based model; forests; *Fuscospora cliffortioides*; southern
34 beech; zone-of-influence.

35 Introduction

36 Individual organisms within natural populations usually vary greatly in size. A
37 description of the distribution of sizes is a common starting point for many de-
38 mographic studies [e.g. 1, 2, 3]. This is especially the case for plants, where size
39 distributions are often considered to convey information regarding the stage of
40 development of a stand or the processes occurring within a population [4]. In
41 the absence of asymmetric competition or size-related mortality, the sizes of indi-
42 viduals within an even-aged cohort should be approximately normally-distributed
43 around a single mode, allowing for some variation in growth rate. More commonly
44 a left-skew is observed during early stages of cohort development; this is attributed
45 to smaller-sized individuals receiving insufficient resources to maintain growth, ul-
46 timately increasing their likelihood of mortality [5, 6]. Size-thinning thereafter
47 reduces the degree of skewness [7, 8, 9] such that the distribution converges on
48 a common maximum size [2]. Finally, as individuals die through disturbance or
49 senescence, and recruitment into lower size classes occurs, populations shift to a
50 size distribution referred to as reverse J-shaped, where a high density of of small
51 individuals is combined with a small number of large dominants. This is a common
52 pattern in old-growth forests, especially those dominated by shade-tolerant species
53 which can persist in small size classes [e.g. 10].

54 A range of statistical models exist to capture these transitions in size distribu-
55 tions [4, 11]. Nevertheless, such simple models are unable to capture the behaviour
56 of many systems. Multimodality of size distributions is widely observed in nature
57 [2, 7, 12]. This is particularly true of plant populations [see Table 1 in 13], even
58 when all individuals are known to have recruited simultaneously [14]. The preva-

59 lence of multimodality is likely to have been underestimated due to a failure to
60 apply appropriate statistical tests [e.g. 15]. In some studies, even when multi-
61 modal distributions are observed, they are overlooked or dismissed as anomalous
62 [e.g. 8, 11, 16].

63 When larger organisms monopolise access to resources it increases the asym-
64 metry of competition among individuals [17]. Small individuals face combined
65 competition from all neighbours larger than themselves, whereas large individuals
66 are unaffected by their smaller neighbours. This is particularly likely to be the
67 case for light competition among vascular plants, where taller stems capture a
68 greater proportion of available radiation and determine access for those beneath
69 [18]. As larger individuals can thereby maintain higher growth rates, incipient
70 bimodality will be reinforced [12], at least until light deprivation causes mortality
71 among smaller individuals [1]. Stand development models are able to generate
72 bimodal patterns when resources for growth become limited [19, 20, 21]. Never-
73 theless, though the potential for bimodality to arise from competitive interactions
74 is well-known, previous models have only been able to reproduce it within a narrow
75 range of parameters [19, 20], leading to the conclusion that it is the least likely
76 cause of bimodality in natural size distributions [12]. A range of alternative mech-
77 anisms might give rise to multimodality, including abiotic heterogeneity whereby
78 large stem sizes are indicative of favourable environmental conditions [22], or se-
79 quential recruitment of overlapping cohorts [12]. Finally, the initial spatial pattern
80 of recruits may itself create complex variation in the sizes of individuals.

81 In this study we argue that instead of being unusual or aberrant, multimodality
82 is an expected outcome whenever strong asymmetries in competition among indi-
83 viduals occur in cohorts of sessile species. We sought to determine the conditions

84 under which multimodal size distributions form in spatially-structured populations
85 using an individual-based modelling approach. Such models have the potential to
86 derive new insights into fundamental ecological processes as they often demonstrate
87 emergent properties which cannot be predicted from population-level approaches
88 [23]. In order to parameterise our models we used a long-term dataset of 250 plots
89 in New Zealand in which the sizes of over 20 000 *Fuscospora cliffortioides* (Hook.
90 f.) Heenan & Smissen (\equiv *Nothofagus solandri* var. *cliffortioides* (Hook. f.) Poole)
91 trees have been monitored since 1974 [9, 24, 25]. These data are used to obtain
92 plausible parameters for our simulation model, which is then employed to explore
93 the causes of multimodality in virtual populations.

94 Our predictions were that (a) the size distribution of individuals would carry
95 a long-term signal of the spatial patterns at establishment, and that (b) asymme-
96 tries in competitive ability would increase the degree of bimodality, which once
97 established would strengthen through time, until resource deprivation removed
98 weaker competitors from the population. Finally, we aimed to test whether (c)
99 manipulating the distance and number of competitors within local neighbourhoods
100 would generate variation in the number and positions of modes within cohort size
101 distributions. Through this work we demonstrate that complex size distributions
102 with multiple modes can be generated within cohorts even in homogeneous envi-
103 ronmental space and when individuals are initially arranged in a regular grid. We
104 show that multimodality is not a transient phase, but is maintained for the pro-
105 jected lifespan of a cohort. Finally, we show that the eventual size reached by any
106 individual depends upon interactions with others in its immediate neighbourhood
107 throughout its lifetime.

108 **Materials and methods**

109 **The simulation model**

110 All parameters used in the text are summarised in Appendix 1. The growth model
111 is derived from a basic energy conservation principle. We assume throughout
112 that resources in the model refer to light (and therefore carbohydrates acquired
113 through photosynthesis), though in principle the model could be extended to other
114 resources with appropriate parameterisation. Recruitment and age-related senes-
115 cence are not included in the model. The resources E that an individual acquires
116 in a unit of time t are distributed between the resources used to increase its size
117 M_g and all other metabolic and maintenance costs M_m . This is expressed math-
118 ematically as a general energy budget $E = M_g + M_m$. Assuming that resource
119 intake scales with biomass m as $E_i \propto m^{3/4}$ [26], and a linear relation between
120 maintenance costs and biomass $M_m \propto m$, we can write a simple individual growth
121 rate equation:

$$\frac{dm}{dt} = am^{3/4} - bm \quad (1)$$

122 where a and b are constants and the units are chosen such that an increase of one
123 unit in biomass requires one unit of resources. A mathematically equivalent model,
124 but with slightly different interpretation, has been proposed previously [18, 27, 28].
125 Equation 1 describes the potential growth rate of an individual in the absence of
126 competition.

127 The potential rate of energy uptake of an individual is reduced when it competes
128 with neighbours and thus they share the available light. In order to take this into

129 account the growth rate in the presence of competition can be expressed as

$$\frac{dm}{dt} = am^{3/4} - bm - \sum_j I(m, m_j, d_j) \quad (2)$$

130 where I_j represents the reduction in biomass growth of a given individual due to
131 competition with another individual j of mass m_j and at a distance d_j from the
132 focal tree. The competitive response is obtained by summing I_j over all interacting
133 neighbours. We only took pairwise interactions into account, summed across all
134 interactions for each individual. This maintained computational efficiency of the
135 simulations [29]. An individual died if its maintenance needs M_m were not met,
136 i.e. if $am^{3/4} - \sum_j I(m, m_j, d_j) < bm$.

137 Spatially explicit interactions among individuals were modelled with a circular
138 zone of influence (ZOI) where A represents the potential two-dimensional space
139 within which a plant acquires resources in the absence of competition. Resource
140 competition between an individual i and its neighbour j is defined as occurring
141 when A_i overlaps with A_j . Within the area of overlap, $A^{(I)}$, resources are dis-
142 tributed among the two individuals, but not necessarily equally. A larger indi-
143 vidual (greater m) will be a stronger competitor, for example by over-topping in
144 light competition, but also potentially through directing greater investment into
145 below-ground resource capture [e.g. 30]. To incorporate asymmetric competition
146 we define $f_m(m, m_j)$ as being the proportion of resources E that an individual of
147 size m obtains from the area within which it interacts with another individual of
148 size m_j . Assuming homogeneous resource intake within A , then E is proportional
149 to $A^{(o)} + f_m(m, m_j)A^{(I)}$, where $A^{(o)}$ is the area within which no interaction occurs
150 ($A - A^{(I)}$).

151 Since in the absence of competition $E = am^{3/4}$, competition will reduce E as
152 follows:

$$E = am^{3/4} - (1 - f_m(m, m_j))A^{(I)} \quad (3)$$

153 and

$$I(m, m_j, d_j) = (1 - f_m(m_j))A_j^{(I)} \quad (4)$$

154 The explicit functional form for asymmetric competition is $f_m(m, m_j) = \frac{m^p}{m^p + m_j^p}$.
155 When $p = 0$ the resources in the zone of overlap are divided equally, irrespective
156 of each individual's size. If $p = 1$ then each individual receives resources in pro-
157 portion to its size, and if $p > 1$ then larger individuals gain a disproportionate
158 benefit given their size. This differs from a previous formulation [31], though their
159 terminology of competitive interactions can be matched to this work as absolute
160 symmetry ($p = 0$), relative symmetry ($p = 1$) and true asymmetry ($p > 1$). The
161 shape of the competition kernel is identical in all cases.

162 This mathematical framework was used to create a spatially-explicit simulation
163 model in which the growth and interactions among large numbers of individuals
164 could be assessed simultaneously.

165 **Model fitting**

166 To obtain realistic parameters for the simulation model we utilised data from
167 monospecific *Fuscospora cliffortioides* forests on the eastern slopes of the Southern
168 Alps, New Zealand. *F. cliffortioides* is a light-demanding species which recruits as
169 cohorts in large canopy gaps, and has a lifespan that seldom exceeds 200 years. The

170 data consisted of records from 20 330 trees situated in 250 permanently marked
171 plots that randomly sample 9 000 ha of forests. Each plot was 20×20 m in size. In
172 the austral summer of 1974–75 all stems >3 cm diameter at breast height (dbh)
173 were tagged and dbh recorded. The plots were recensused during the austral
174 summers of 1983–84 and 1993–94. Only stems present in more than one census
175 were included. Previous work on this system has confirmed a dominant role for
176 light competition in forest dynamics [9, 18].

177 We tested the tree size distribution from the first survey of each plot for multi-
178 modality by fitting a finite mixture model of one, two and three normal distribu-
179 tions (see Appendix 1). We employed an expectation-maximisation (EM) algorithm
180 [32] within the R package `FlexMix` 2.3-4 [33] and utilised the Bayesian Informa-
181 tion Criterion (BIC) to decide whether each size distribution was unimodally or
182 multimodally distributed.

183 In order to fit the simulation model to the data we estimated the mass m of the
184 trees by allometric relation $\text{dbh} = C_{\text{dbh}} m^{3/8}$ [26, 34], where C_{dbh} was taken as a free
185 parameter. The area A of the circle representing the potential space for resource
186 acquisition was given by $cA = am^{3/4}$ where c is a proportionality constant. A
187 linear relation between dbh and radius of the zone of influence was chosen, and
188 a high degree of asymmetric competition ($p = 10$). The latter improved overall
189 fit of the models, indicating a role for asymmetric competition in driving stand
190 dynamics.

191 For each of 250 plots we began the simulation model with the observed stem
192 sizes from 1974 attached to points randomly distributed in space. The simulation
193 was run for 19 model years, developed in time increments δt which nominally
194 correspond to 10 weeks (for simplicity there is no seasonality of growth). An

195 individual's growth is given by:

$$\delta m_i = \left[am_i^{3/4} - bm_i - \sum_j \frac{m_j^p}{m_i^p + m_j^p} cA_j^{(I)} \right] \delta t \quad (5)$$

196 In each Monte Carlo iteration individuals m_i were selected at random and their
197 size updated. A search algorithm was employed to find values of a and b which gave
198 the best fit to the observed individual growth rates with Pearson's χ^2 , averaged
199 across the ensemble of simulations. Note that the model was fit to the growth
200 rates of individual stems based on repeated measurements, rather than stand-level
201 properties such as size distributions.

202 Having obtained suitable values for a and b we performed simulations to com-
203 pare the size distributions as predicted by the model (assuming initially random
204 stem positions) with the empirical distributions observed in the data set. These
205 were initiated using size distributions from stands in the *F. cliffortioides* dataset
206 in which the mean stem diameter was small ($\bar{d} < 15$ cm), then run until the mean
207 reached a medium ($15 \text{ cm} \leq \bar{d} < 22$ cm) or large ($\bar{d} \geq 22$ cm) stem size. Estimates
208 of size-dependent mortality rate were also obtained and compared with empirical
209 outputs as in [9]. This provides an independent evaluation of model performance
210 as mortality rates were not used to parameterise the model.

211 Exploring multimodality in size structure

212 The simulator with fitted parameters as described above was used to explore the
213 factors which cause multimodal size distributions to form. We tracked the devel-
214 opment of size structures in simulated stands with differing initial spatial patterns
215 and symmetry of competition. In these simulations all individuals were of identical

216 initial size.

217 First 2100 spatial patterns were generated, each containing a distribution of
218 points with x and y co-ordinates in a virtual plot of 20×20 m. Equal numbers
219 patterns were clustered, random and dispersed. Random patterns were produced
220 using a uniform Poisson process with intensity $\lambda = 0.05$ points m^{-2} . Clustered
221 patterns were created using the Thomas process. This generated a uniform Poisson
222 point process of cluster centres with intensity $\lambda = 0.005$. Each parent point was
223 then replaced by a random cluster of points, the number of points per cluster being
224 Poisson-distributed with a mean of 10, and their positions as isotropic Gaussian
225 displacements within $\sigma = 1$ from the cluster centre. Dispersed patterns were
226 produced using the Matern Model II inhibition process. First a uniform Poisson
227 point process of initial points was generated with intensity $\lambda = 0.06$. Each initial
228 point was randomly assigned a number uniformly distributed in $[0,1]$ representing
229 an arrival time. The pattern was then thinned by deletion of any point which
230 lay within a radius of 1.5 units of another point with an earlier arrival time.
231 All patterns were generated in R using the `spatstat` package [35]. Each pattern
232 contained roughly 500 points (clustered $N = 501 \pm 2.7$, random $N = 501 \pm 0.8$,
233 dispersed $N = 488 \pm 0.7$). The slightly lower number of points in the dispersed
234 pattern reflects the inherent difficulties in generating a dense pattern with a highly-
235 dispersed structure and has no qualitative effect on later analyses. Although the
236 density within starting patterns was approximately a quarter of that observed in
237 the empirical data, initial density has a limited effect on final outcomes since its
238 main effect is to reduce the time until points begin to interact [36], and lower point
239 densities increased computational speed, allowing for greater replication.

240 A number of further patterns were generated to explore the influence of specific

241 parameters. First, a regular square grid was used with a fixed distance of 1.5 or
 242 3 m between individuals. Next, groups of individuals were created in which all
 243 individuals within groups were 3 m apart, but with sufficient distance among
 244 groups that no cross-group interactions could take place. Groups contained either
 245 two individuals (pairs), three individuals in a triangular arrangement (triads) or
 246 four individuals in a square arrangement (tetrads). The total starting population
 247 in each pattern was approximately 7500 individuals.

248 We ran simulations of the spatially-explicit individual-based model, varying
 249 the degree of asymmetric competition p . The points generated above became
 250 individual trees represented as circles growing in two-dimensional space. Each
 251 individual was characterised by its mass m and co-ordinates. In order to model
 252 mortality, an individual was removed from the simulation if carbon losses exceeded
 253 gains, that is if $[am_i^{3/4} - bm_i - \sum_j \frac{m_j^p}{m_i^p + m_j^p} A_j^{(I)}] < 0$.

254 The predicted size distribution and mortality rate of clumped, random and
 255 dispersed starting patterns were obtained from ensemble averages of 700 simula-
 256 tions corresponding to the point processes generated above. m was a continuous
 257 variable but in order to derive the size distribution, growth and death rates we cal-
 258 culated size frequencies based on 10 kg biomass bins. Since the death rate changes
 259 through time due to alterations in the size structure of the community, we present
 260 the average death rate for each size class across all time steps in simulations, which
 261 run for 460 model years (at which point only a few very large stems remain). This
 262 allows sufficient resolution for figures to be presented as effectively continuous re-
 263 sponses rather than histograms, and is equivalent to a landscape-scale aggregation
 264 of size-dependent mortality data across a series of stands of differing ages.

265 Results

266 Analysis of the New Zealand forest plot dataset revealed multimodal distributions
267 in 179 plots in 1974, 163 plots in 1984 and 152 plots in 1993 from of a total of 250
268 plots in each survey. This represents 66% of plots, showing that multimodality is
269 more common than unimodality within these forests (see Appendix 2).

270 The simulation model was fit to the observed individual growth rates in the
271 *F. cliffortioides* dataset and provided a robust representation of the empirically-
272 measured patterns. The fitted parameters (a , b and C_{dbh}) are shown in Appendix
273 1. The effectiveness of the model was assessed through its ability to capture size-
274 dependent mortality rates, which were an emergent property of the system and
275 not part of the fitting process. Size distributions thus obtained were qualitatively
276 similar to those observed in the empirical dataset [9]; see Appendix 3.

277 Subsequent simulation modelling used the parameters derived from the *F. clif-*
278 *fortioides* dataset (a , b , C_{dbh}) and created simulated forests to investigate the
279 potential origins of multimodal patterns. Using stochastically-generated starting
280 patterns, major differences were evident in the patterns of growth and survival
281 depending on the degree of competitive asymmetry p and the initial spatial con-
282 figuration (Fig. 1).

283 With completely symmetric competition among individuals ($p = 0$), average
284 tree growth in clustered patterns was greater than in either random or dispersed
285 patterns (Fig. 1a). This unexpected result can be attributed to the high rate of
286 density-dependent mortality in very early time steps (Fig. 1d). Initial mortality in
287 random patterns reduced the population to be comparable with dispersed patterns,
288 compensating for the slight initial differences in abundance. Clustered populations

289 remained larger in average stem size (Fig. 1a) as the result of a smaller final
290 population size (Fig. 1d), an effect which developed rapidly and was maintained
291 beyond the plausible 200-year lifespan of *F. cliffortioides*.

292 In the absence of asymmetric competition ($p = 0$), starting patterns had a
293 limited effect on final size distributions, with only minor increases in skewness
294 in clustered populations at advanced stages of development (Appendix 4). In all
295 cases size distributions remained unimodal. It is therefore apparent that varia-
296 tion in initial spatial patterns is not in itself sufficient to generate multimodality
297 in size distributions, at least not unless the average distance among individuals
298 exceeds their range of interaction, which is highly unlikely in the context of plant
299 populations.

300 The introduction of weak asymmetry ($p = 1$) tended to increase the mean size
301 of individuals while causing reductions in population size (Fig. 1b,e) and dimin-
302 ishing the differences among initial patterns, such that with strong asymmetry
303 ($p = 10$) the differences in final size between starting patterns were negligible
304 (Fig. 1c). Strong asymmetry also caused population sizes to converge within the
305 likely lifespan of the trees, irrespective of starting conditions, and at a lower fi-
306 nal level (Fig. 1f). Reduced differences among initial patterns with increasing
307 asymmetry arose because fewer small trees survived around the largest tree in the
308 vicinity, which caused patterns to converge on a state with dispersed large indi-
309 viduals and smaller individuals in the interstices. More left-skewed distributions
310 also emerged as a consequence of the low tolerance of individuals to depletion
311 of resources (individuals failing to obtain sufficient resources for their metabolic
312 needs died immediately). Thus the small individuals die soon after their resource
313 acquisition area is covered by the interaction range of a larger individual. Such

314 left skew would be reduced for species capable of surviving long periods of time
315 with low resources either through tolerance or energy reserves.

316 Increasing competitive asymmetries caused size distributions to exhibit slight
317 multimodality with a lower frequency of individuals in the smaller size class at 150
318 years (Appendix 5). Given entirely random starting patterns, more pronounced
319 bimodality emerged as the degree of asymmetric competition increased. Further-
320 more, the model predicted a U-shaped size-dependent mortality rate, qualitatively
321 consistent with a pattern in the empirical data (Appendix 6; compare Fig. 5 in
322 [9]). This trend intensified with increasing asymmetric competition, and was ab-
323 sent when resource division was symmetric. It occurred because in large trees
324 the majority of resources are required for maintenance, and therefore even a rel-
325 atively small amount of competition ultimately increases their mortality rate. In
326 the absence of asymmetric competition the death rate of large trees approached
327 zero.

328 Greater insights into the causes of multimodality are revealed through the use
329 of designed spatial patterns in which the timing of interactions within model devel-
330 opment can be precisely controlled. These illustrate that the separation between
331 modes is determined by the distance among competing individuals under asym-
332 metric competition (Fig. 2 and Appendix 7). The size structure can therefore
333 provide an indication of the dominant distance over which individuals are compet-
334 ing, though separation of modes will be less clear when a strict grid is absent. Note
335 that the position of the right-hand mode remains identical, and it is only the mode
336 of the subordinate individuals which shifts to a smaller size class. Highly-dispersed
337 patterns give rise to more complex size distributions through their development
338 when asymmetric competition is present. In the most extreme case, when initial

339 patterns are gridded, each individual interacts with a series of neighbours as its
340 size increases, leading to a complex multimodal pattern, at least until continued
341 mortality removes smaller size classes (Fig. 3). Note that the modes are more
342 clearly distinguished than is the case for random starting patterns where distances
343 among individuals vary (compare Fig. 8c).

344 The patterns generated by small groups of interacting individuals at equal dis-
345 tances apart with asymmetric competition lead to size distributions with a number
346 of modes equal to the number of individuals within each group. For patterns de-
347 rived from pairs of individuals, the size distribution is bimodal, and in similar
348 fashion triads and tetrads produce size distributions with three and four modes
349 respectively (Fig. 4). Each mode corresponds to the discrete ranking of individuals
350 within groups. This indicates that in gridded populations, as might be observed
351 in plantations or designed experiments, the number of modes is determined by the
352 effective number of competitors.

353 Discussion

354 Multimodality in cohort size distributions is the outcome, rather than the cause,
355 of asymmetric competition among individuals of varying size. Regardless of ini-
356 tial small-scale starting patterns, size distributions remain unimodal in the case
357 of symmetric competition among individuals. Only when larger individuals are
358 able to acquire a greater proportion of resources from shared space does bimodal-
359 ity begin to emerge. Spatial patterns of established individuals can modulate
360 these interactions, with complex multimodal distributions generated when indi-
361 viduals are either regularly or highly dispersed in space. The number of modes

362 corresponds to the number of effective competitors and their separation is a con-
363 sequence of average distances among individuals. Note that our simulations do
364 not incorporate continuous recruitment; this is a reasonable assumption for sys-
365 tems such as *F. cliffortioides* forests, where large-scale disturbances are followed
366 by stand replacement.

367 Asymmetric competition will lead to multimodal distributions at some point
368 during stand development. We extend upon previous studies [e.g. 37] by provid-
369 ing a general framework for predicting and interpreting complex size distributions
370 in spatially-structured and even-aged populations. Under light competition the
371 modes will correspond to discrete and well-defined canopy layers. In [13] a se-
372 ries of controlled experiments were conducted to investigate size distributions in
373 populations of annual plants, finding in many cases that distributions with two or
374 three modes were observed. Our results allow for a fuller interpretation of these
375 earlier findings, as we have shown that the number of modes reflects the number
376 of effective competitors, placing a limit on the complexity of size distributions.
377 As demonstrated in Figs. 2 and 4, the larger mode remains in the same position
378 regardless of the size at which competition begins. This highlights that those in-
379 dividuals in larger size classes are almost unaffected by competition during stand
380 development.

381 Even when all individuals in a cohort begin with identical size, small fluctua-
382 tions in the acquisition of shared resources lead to a multimodal size distribution,
383 regardless of whether the initial pattern was random, dispersed or clustered. The
384 size distribution is not affected by differences in the initial spatial structure at small
385 scales due to the death of close neighbours early in stand development. A similar
386 result was found by [36], who argue that the importance of recruitment patterns in

387 generating asymmetries in competition may have been over-stated. Likewise initial
388 density will have a limited effect on final size distributions as its main influence is
389 on the time at which individuals begin to interact [36]. Therefore, while local in-
390 teractions undoubtedly do cause competitive asymmetries [e.g. 17], these are more
391 relevant in determining the pattern of mortality during self-thinning rather than
392 final size distributions, so long as the distances over which competition influences
393 growth are larger than the characteristic scales at which initial spatial structuring
394 occurs. In dense aggregations of recruiting plants this is likely to be the case.

395 The model implies only a single resource for which individuals compete. It is
396 typically assumed that above-ground competition for light is asymmetric, whereas
397 below-ground resources are competed for symmetrically [38], though the latter
398 assumption may not always be true [e.g. 39, 40]. More complex zone-of-influence
399 models can take into account multiple resources and adaptive allometric changes on
400 the part of plants in response to resource conditions [e.g. 41, 42]. Indeed, plasticity
401 can diminish the impact of asymmetric competition [41, 43]. Although below-
402 ground interactions are challenging to measure directly, there is good evidence
403 that above- and below-ground biomass scale isometrically [44] which justifies the
404 use of above-ground biomass to infer potential root competition. Previous work
405 using the same data has identified a dominant role for light competition among
406 smaller stems, with nutrient competition important at all stem sizes [18].

407 Forest mensuration tends to overlook the shape of size distributions in favour
408 of summary statistics [e.g. mean size, coefficient of variation, maximum size; 45]
409 and may therefore miss out on valuable contextual information. While the utility
410 of size distributions as a predictive tool for modelling dynamics has been fre-
411 quently overstated [see 46], they can nonetheless remain a valuable indicator of

412 past dynamics. One outcome of bimodality arising from asymmetric competition
413 is that large and small individuals have differing spatial patterns, with the larger
414 dispersed in space and the smaller confined to the interstices generated by the
415 dominant competitors [47]. This can be used as a diagnostic tool as it allows this
416 mechanism to be distinguished from abiotic heterogeneity, leading to clustering
417 of similar sizes, or independent sequential recruitment, leading to a lack of co-
418 associations between size classes [12]. Likewise in mixed-species stands succession
419 can cause a multimodal pattern to emerge through aggregation of several unimodal
420 cohorts, persisting throughout stand development [10]. The interplay between size
421 distributions, plant traits and disturbance can generate complex emergent pat-
422 terns in forest dynamics at the landscape scale [48]. Bimodality generated by size
423 competition among individuals is a distinct phenomenon from the bimodality in
424 inherited size across species which is often observed in mixed-species communities
425 [e.g. 49]. Where size histograms combine individuals from multiple species, the
426 causes of bimodality are likely to include long-term evolutionary dynamics in ad-
427 dition to direct competition among individuals. Contextual information on spatial
428 patterns, disturbance regimes and community composition are therefore essential
429 to interpreting size distributions in natural systems.

430 Our models are based upon parameters obtained from a long-term dataset and
431 can therefore be immediately transferred to a predictive framework. While the
432 exact terms are most suited to the *Fuscospora cliffortioides* forests which form
433 the basis of this work, it is likely that they will be applicable to any monospecific
434 plant population. Bimodal size distributions might be overlooked where aggregate
435 curves are drawn as composites of a large number of plots, which will tend to
436 average out differences, or where appropriate statistical tests are not employed.

437 We find that 66% of plot size distributions in our data are bimodal. It is likely
438 that these do not all represent single cohorts; for example, a severe storm in 1972
439 opened the canopy in some plots and allowed a recruitment pulse [24, 50]. Irre-
440 spective of this, our growth model is able to capture subsequent stand development
441 regardless of the origin of the bimodality (see Appendix 3). Our results also show
442 that multimodality can act as an indicator of asymmetric competition. Thomas
443 & Weiner [31] present evidence that the degree of asymmetry in natural plant
444 populations is strong, with larger individuals receiving a disproportionate share of
445 the resources for which they compete ($p \gg 1$). The phenomenon of multimodality
446 should therefore be widespread.

447 In conclusion, and in contrast with a previous review of bimodality in cohort
448 size distributions [12], we contend that asymmetric competition is the leading can-
449 didate for explaining multimodal size distributions, and is its cause rather than the
450 outcome. Previous simulation results suggesting that the parameter space within
451 which multimodality occurs is limited were based on stand-level models. Through
452 the use of individual-based models it can be demonstrated that multimodality is
453 an expected outcome for any system in which larger individuals are able to control
454 access to resources, and where individuals compete in space. The strength of these
455 asymmetries determines the degree to which multimodality is exhibited, while the
456 number and separation of modes are determined by the number of effectively-
457 competing individuals and the distances among them. While multimodality may
458 be a transient phase within the development of our models, many forest stands
459 exhibit non-equilibrium conditions, and indeed most natural plant populations are
460 prevented by intermittent disturbance from advancing beyond this stage [24, 50].
461 Consistently unimodal size distributions should be seen as the exception rather

462 than the rule.

463 **Data accessibility**

464 Data were obtained from New Zealand's National Vegetation Survey Databank and
465 can be accessed at [http://datastore.landcareresearch.co.nz/dataset/multimodal-](http://datastore.landcareresearch.co.nz/dataset/multimodal-size-distributions-in-spatially-structured-populations)
466 [size-distributions-in-spatially-structured-populations](http://datastore.landcareresearch.co.nz/dataset/multimodal-size-distributions-in-spatially-structured-populations). All C code used to run the
467 simulations can be obtained from <https://github.com/jorgevc/IMB-SizeDependent>.

468 **Competing interests**

469 We have no competing interests.

470 **Authors' contributions**

471 MPE and JV conceived and designed the study; RBA and DAC provided data;
472 JV carried out the statistical analyses under guidance from MPE, DAC and RBA;
473 JV and MPE prepared the first draft of the manuscript. All authors contributed
474 towards manuscript revisions and gave final approval for publication.

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622 **Figure captions**

623 **Figure 1.** Cohort-level characteristics of stands with either random, clustered or
624 dispersed initial starting patterns over t years (simulation time). (a–c) Mean tree
625 size in kg with increasing levels of asymmetric competition p (0, 1, 10), note that
626 (a) has a reduced y -axis length; (d–f) mean number of surviving individuals N per
627 20×20 m plot with competition varying from symmetric ($p = 0$) to weakly ($p = 1$)
628 and strongly asymmetric ($p = 10$). Each line is derived from an ensemble average
629 of 700 simulations.

630

631 **Figure 2.** Separation between modes with varying distance of competing neigh-
632 bours and strong asymmetric competition ($p = 10$). Size distributions of stands
633 composed by pairs of equidistant individuals after 200 years of development. Solid
634 line: individuals spaced at 1.5 m, dashed line: individuals spaced at 3 m. Each
635 line is derived from an ensemble average of 700 simulations.

636

637 **Figure 3.** Emergent size distribution through stand development given an initially
638 gridded starting pattern. Individuals separated by 1.5 m from their neighbors and
639 with strong asymmetric competition ($p = 10$). Panels show distribution at 150,
640 200, 230 and 250 years. Each plot is derived from an ensemble average of 700
641 Monte Carlo simulations.

642

643 **Figure 4.** Size distributions of stands composed of groups of two, three and four
644 equidistant competing individuals (pairs, triads and tetrads respectively) with 3
645 m of separation among individuals in each group. Asymmetric competition set

646 at $p = 10$. Each line is derived from an ensemble average of 700 simulations and
647 shows the distribution at 250 years.

Figure 1: Cohort-level characteristics of stands with either random, clustered or dispersed initial starting patterns over t years (simulation time). (a–c) Mean tree size in kg with increasing levels of asymmetric competition p (0, 1, 10), note that (a) has a reduced y -axis length; (d–f) mean number of surviving individuals N per 20×20 m plot with competition varying from symmetric ($p = 0$) to weakly ($p = 1$) and strongly asymmetric ($p = 10$). Each line is derived from an ensemble average of 700 simulations.

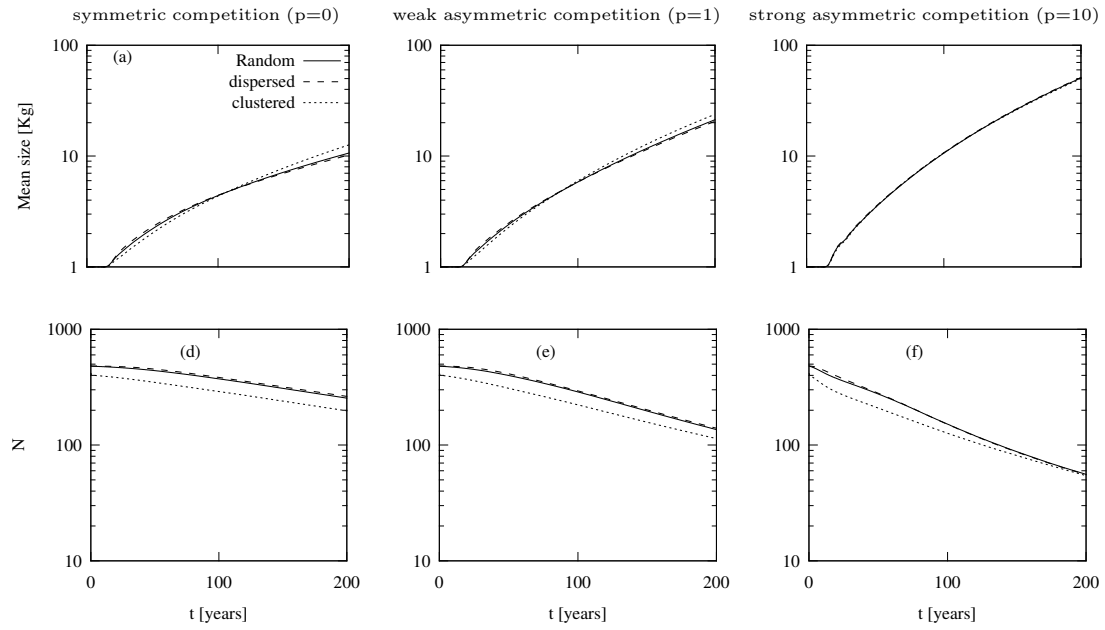


Figure 2: Separation between modes with varying distance of competing neighbours and strong asymmetric competition ($p = 10$). Size distributions of stands composed by pairs of equidistant individuals after 200 years of development. Solid line: individuals spaced at 1.5 m, dashed line: individuals spaced at 3 m. Each line is derived from an ensemble average of 700 simulations.

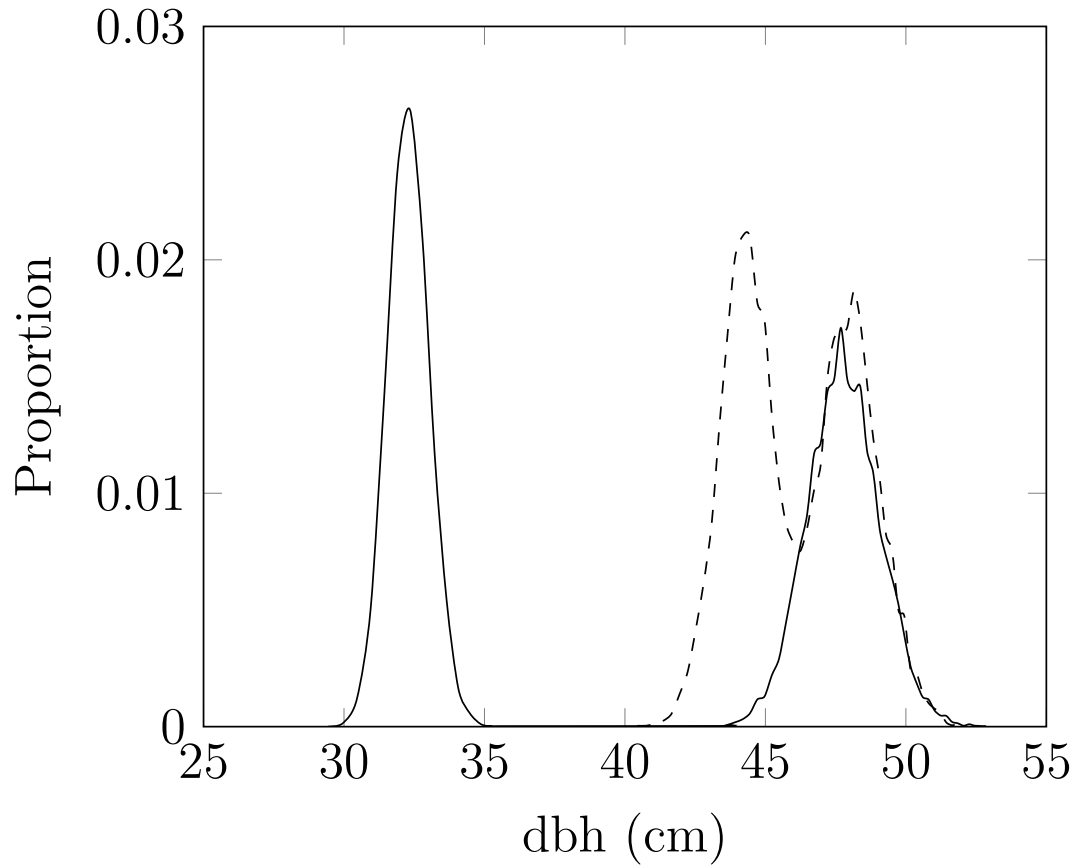


Figure 3: Emergent size distribution through stand development given an initially gridded starting pattern. Individuals separated by 1.5 m from their neighbors and with strong asymmetric competition ($p = 10$). Panels show distribution at 150, 200, 230 and 250 years. Each plot is derived from an ensemble average of 700 Monte Carlo simulations.

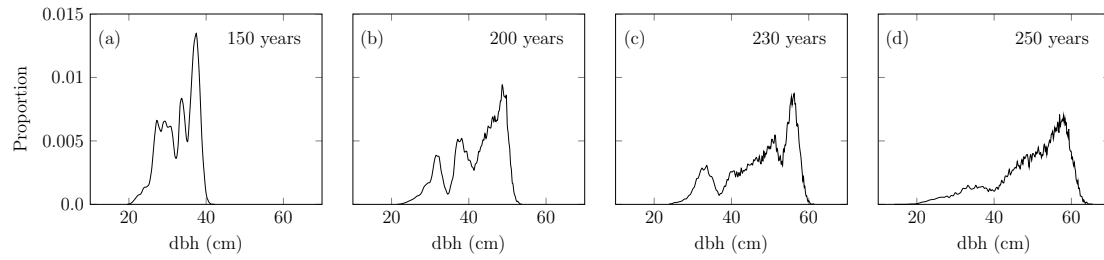
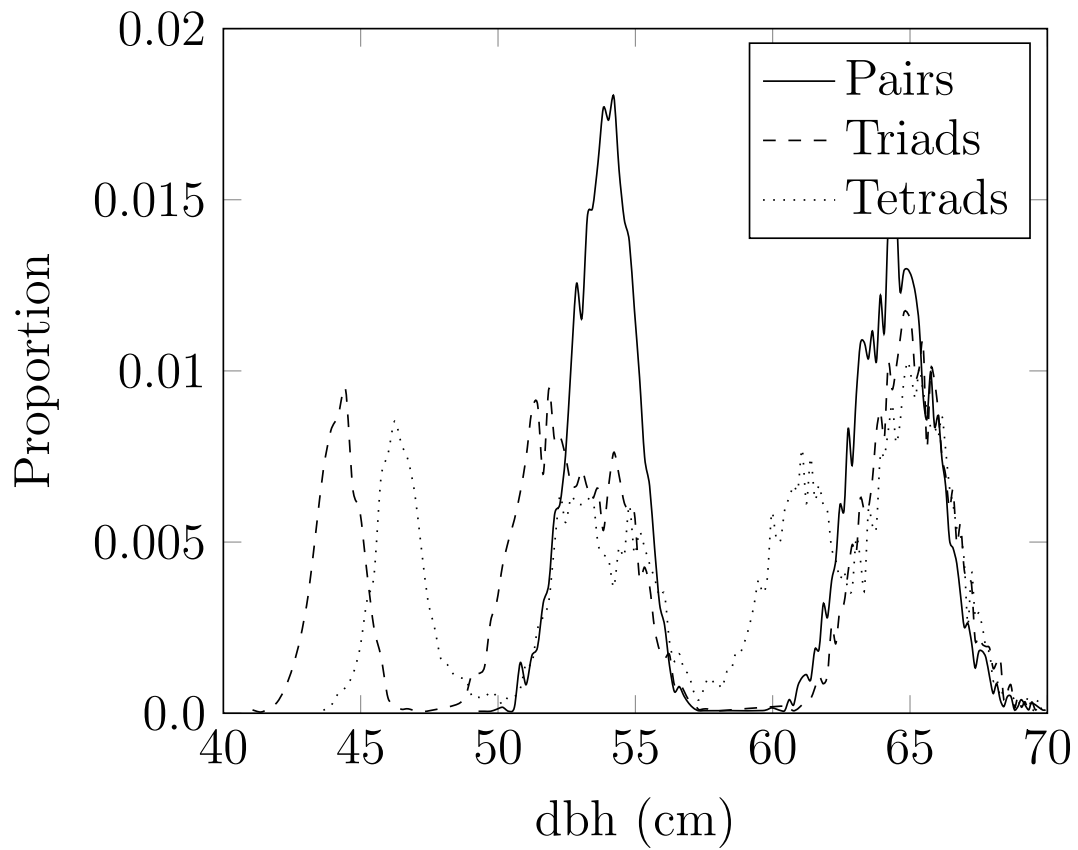


Figure 4: Size distributions of stands composed of groups of two, three and four equidistant competing individuals (pairs, triads and tetrads respectively) with 3 m of separation among individuals in each group. Asymmetric competition set at $p = 10$. Each line is derived from an ensemble average of 700 simulations and shows the distribution at 250 years.



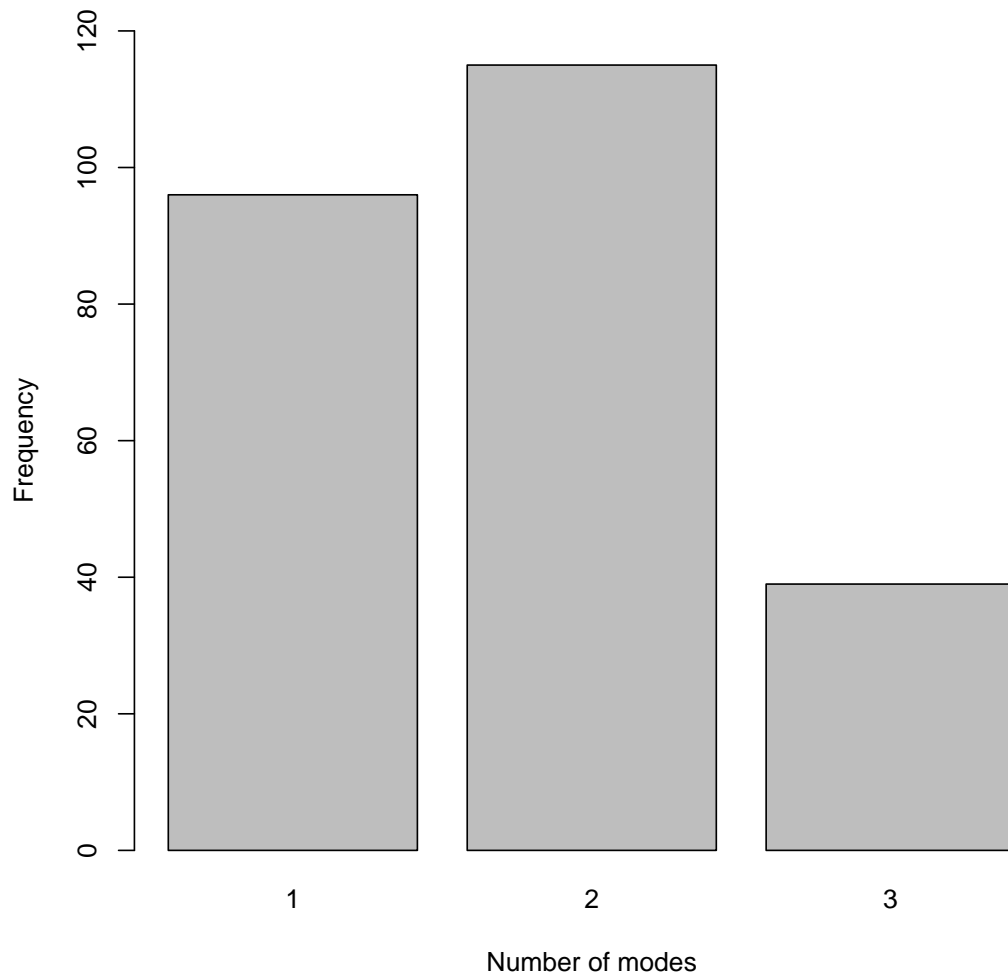
Appendix 1

Table 1: Model terms as used in the text, separated between fitted parameters obtained from field data and free variables at the individual and stand level.

Symbol	Value	Units	Definition
Fitted parameters			
a	2.5×10^{-3}	$10 \times \text{kg}^{-3/4} \text{year}^{-1}$	Conversion factor between $m^{frac-3/4}$ and E
b	2.5×10^{-4}	$10 \times \text{kg}^{-1}$	Resource cost for maintenance per unit biomass
C_{dbh}	9.4	$\text{cm}/10 \times \text{kg}^{3/8}$	Allometric relation between biomass and dbh
Individual-level parameters			
m	variable	$10 \times \text{kg}$	Biomass of an individual
d_j	variable	m	Distance of an individual i to its neighbour j
A_j^I		m^2	Area of interaction between an individual i and its neighbour j
Stand-level parameters			
p	fixed	dimensionless	Degree of competitive asymmetry. $p = 0$ corresponds to symmetric competition while $p > 0$ indicates asymmetric competition
E	equation (3)	$10 \times \text{kg} \text{year}^{-1}$	Resource intake rate of an individual
$I(m, m_j, d_j)$	equation (4)	Resource year^{-1}	Reduction of resource intake rate due to competition
$f_m(m, m_j)$	$\frac{m^p}{m^p + m_j^p}$	dimensionless	Fraction of resources that an individual of biomass m obtains from the area of interaction with an individual of biomass m'

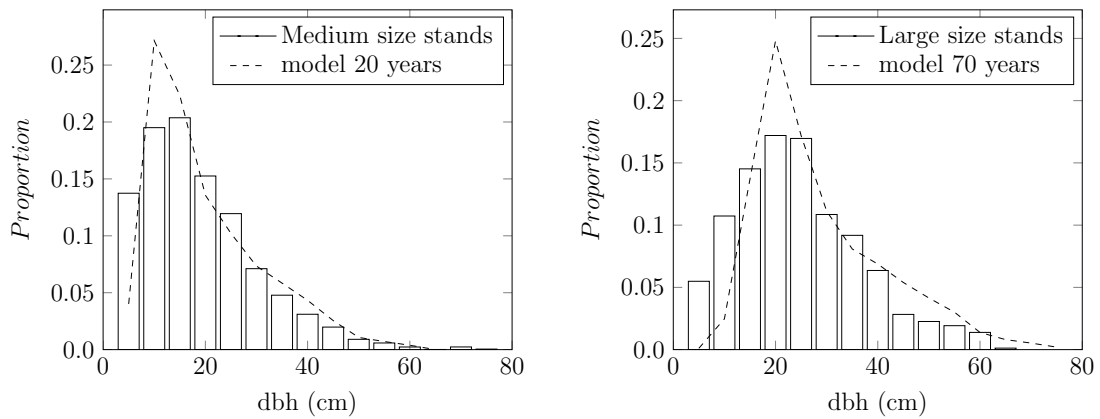
649 **Appendix 2**

Figure 5: Frequency of *Fuscospora cliffortioides* plots in New Zealand exhibiting uni- or multimodality in size distribution as determined by finite mixture models testing for the presence of one, two or three modes. Data from initial 1974–1975 surveys.



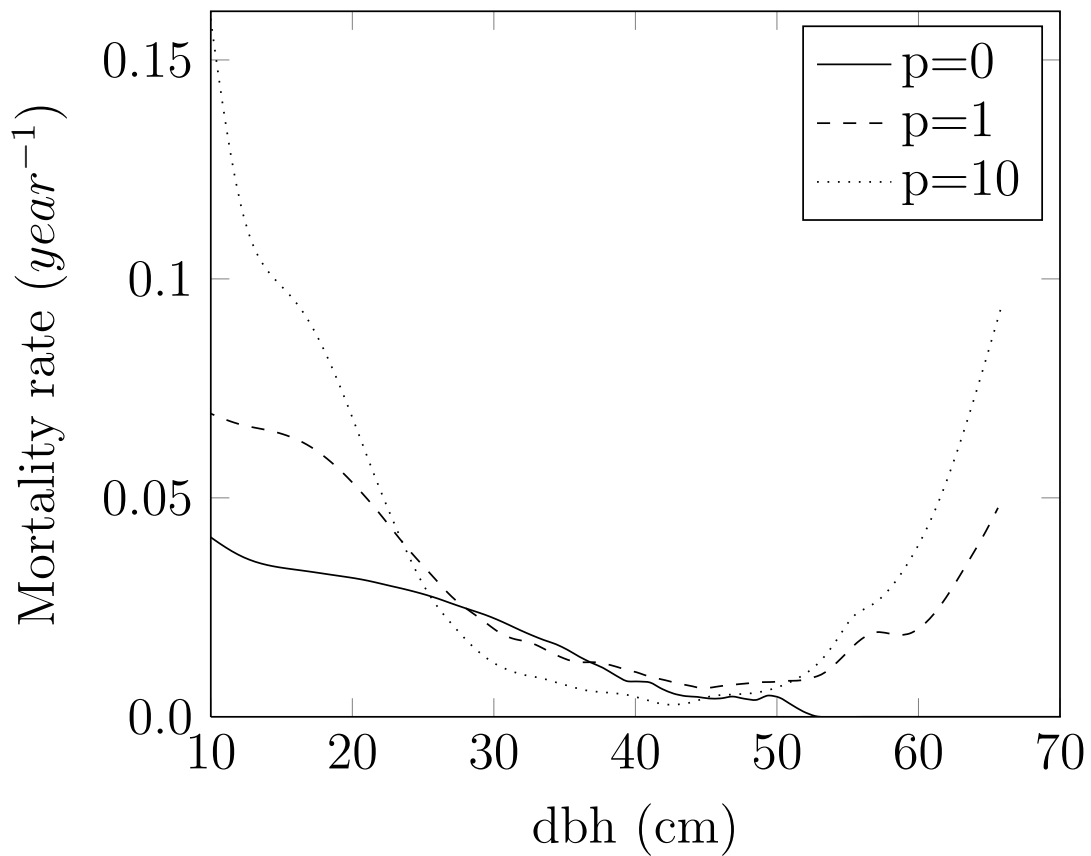
650 **Appendix 3**

Figure 6: Comparative figure to match Fig. 3 in Coomes & Allen 2007. Histograms show distributions of diameter at breast height (dbh; cm) of stands in which mean stem sizes were of medium (15–22 cm dbh; a) and large mean size (>20 cm dbh; b) in 1974. Simulations began with trees in random positions following a size distribution taken from the 117 stands with small mean stem size (<15 cm) in 1974. Dashed lines indicate patterns in simulated stands after 20 or 70 years of model time respectively. This is the ensemble average of $117 \times 4 = 468$ simulations.



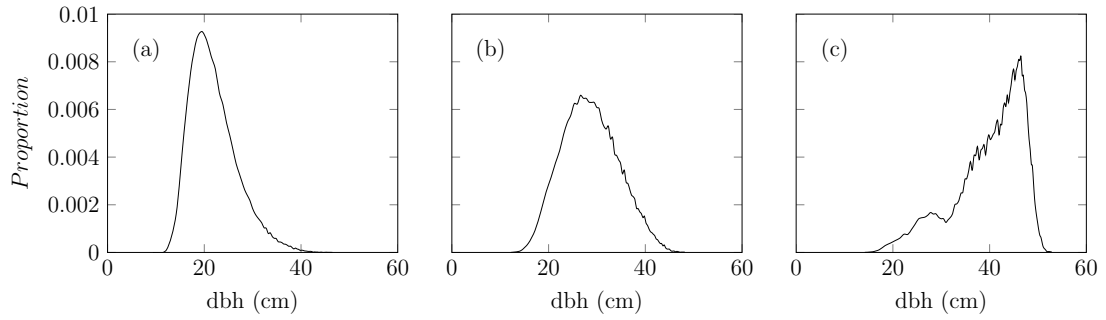
651 **Appendix 4**

Figure 7: Mortality rate as a function of tree size. Solid line for symmetric competition, dashed and dotted lines correspond to increasing asymmetric competition. Derived from an ensemble average of 700 simulations, each of which is run for a nominal 460 years, and showing the cumulative function over the whole time period.



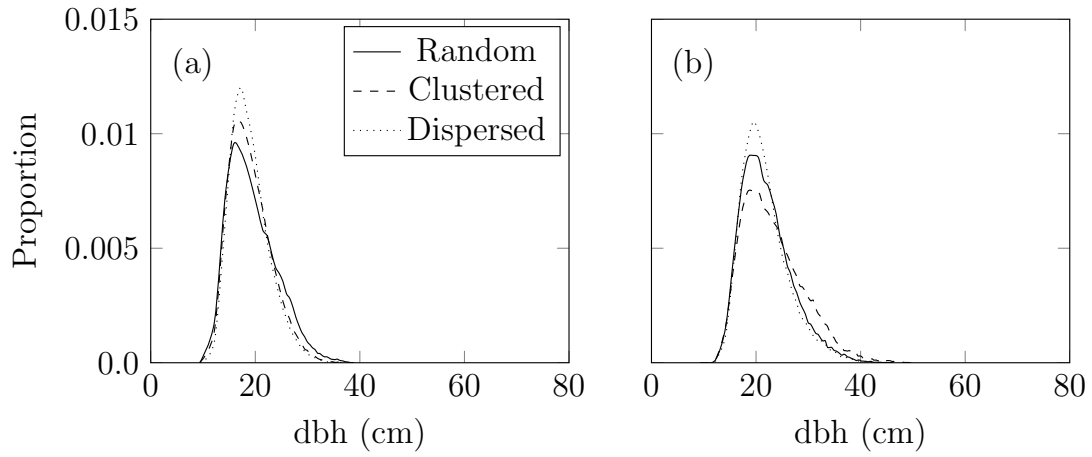
652 **Appendix 5**

Figure 8: Size-frequency histograms for simulated stands. All plots represent 150 years of stand development with increasing levels of asymmetric competition p (0, 1, 10) and random initial pattern. Each plot is derived from an ensemble average of 700 simulations.



653 **Appendix 6**

Figure 9: Size distributions of populations with symmetric competition among individuals ($p = 0$) but variation in initial pattern (random, dispersed, clustered). Panels show distribution at 150 and 200 years. Each plot is derived from an ensemble average of 700 simulations.



654 **Appendix 7**

Figure 10: Effect of increasing distance between paired individuals within simulations (as Fig. 2) on separation between modes in the emergent size distribution. Note that increasing distance reduces the separation of modes by increasing the model time required for two individuals to begin competing for resources.

