

Fig. 1. Dependence of persistence measure on defining parameters T and W : calculated persistence rates as a function of (A) time T after fitting for various values of the time-window length W and (B) W for $T = 4$ years.

width W of the time window for observing battery orders, and the stability of the measure with regard to variations of these parameters must be understood.

For our dataset, the persistence measure is only meaningful for $T \geq W$ [for $T < W$, every patient has a battery order within W because the fitting date is defined as being equal to the date of the first battery order after HA order (Saunders et al. 2020)]. Figure 1A indicates that after $T = 2$ years, the persistence rate stays almost constant as a function of T , that is, the proportion of persistent HA users remains the same over time. Therefore, analyses will not be affected by the choice of T if $T \geq 2$ years. However, over time the sample size is reduced, thus we use $T = 2$ years as a default value.

The dependence on W is necessarily more pronounced, since for $W \rightarrow 0$, the time window for observing battery orders shrinks, so the persistence measure must approach zero. Figure 1A shows that the calculated persistence rate indeed varies with W . For consistency, it is thus important to keep W fixed across all analyses within a project.

For our analyses we use $W = 18$ months because this value falls at the cross-over from a fast to a slow increase in the persistence rate (Fig. 1B). This value of W implies that patients may remain classified as “persistent” if they use their HAs for at least 1/3 of the time (a nominal 6-month supply lasting 18 months).

Thus, in the analyses presented below, we choose $T = 2$ years and $W = 18$ months.

RESULTS: VALIDITY OF PERSISTENCE MEASURE

Connection Between HA Use Persistence and HA Appointments

It is plausible that there will be an association between HA use persistence and the time span over which HA appointments occur after fitting. This is corroborated by Figure 2, which shows a clear difference between nonpersistent and persistent users in the distribution of time intervals between HA fitting and the last recorded HA appointment. In fact, for 32.3% of nonpersistent users, their last appointment coincides with the HA fitting.

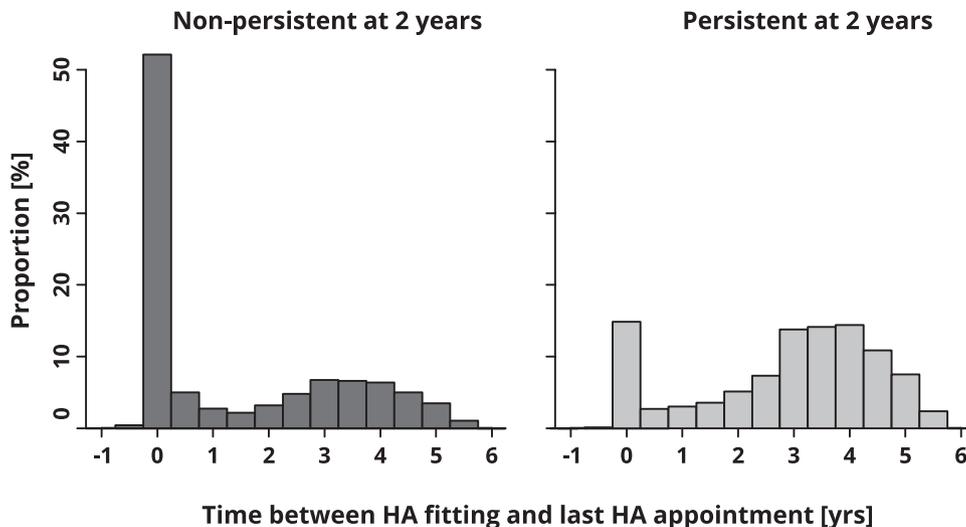


Fig. 2. Distribution of time intervals between HA fitting and last recorded HA appointment for patients nonpersistent and persistent at 2 years after HA fitting. Negative time intervals occur if the last HA appointment found in the outpatient records took place before the first battery order after HA order (which is used as proxy for HA fitting) and thus are artifacts. HA, hearing aid.

Connection Between HA Use Persistence and Short-Term HA Usage (IOI-HA Question 1)

Figure 3A shows the daily reported HA use from the IOI-HA for persistent and nonpersistent patient groups, respectively. It is evident that persistent users report greater daily HA use early on after fitting. However, this should not be interpreted to suggest that more HA use shortly after fitting results in greater long-term persistence, because the consequence of low daily HA use is low battery usage and thus a higher likelihood of being misclassified as nonpersistent by our measure. A better interpretation of Figure 3A might be that daily HA use tends to remain consistent over time. Regardless, the strong association supports the validity of our persistence measure. The patterns seen in Figure 3A remain similar when plotted for subsamples according to the time of IOI-HA report relative to fitting, that is, they do not appear to depend on when the IOI-HA was submitted.

Dependence of Persistence on Age and Pure-Tone Average

Figure 3B illustrates that persistence is dependent on age and four-frequency pure-tone average (4F-PTA; bilateral mean of thresholds at 0.5, 1.0, 2.0, and 4.0 kHz) such that persistence increases with hearing loss (up to around 80 dB HL) and is highest for individuals aged 70–80 years. The former finding is consistent with the explanation that HAs offer increasing benefit up to a certain level of hearing loss (i.e., 80 dB). The latter likely reflects decreasing stigma as age increases, which, however, is counterbalanced by decline in cognitive and physical function beyond about 75 years. Persistence then declines most rapidly among those with milder hearing loss. These results are in line with other work (Gopinath et al. 2011; Bainbridge & Ramachandran 2014), thus providing further evidence for the validity of the persistence measure.

DISCUSSION

The rationale underlying our HA persistence measure is that HA use can be inferred from battery-order history and, more specifically, that patients maintain HA use for a certain time period after a battery order. We provide evidence for the validity of the measure through its association with HA outpatient appointments, HA short-term usage (via IOI-HA) and patient age and PTA. The measure is robust regarding variations of T , but the more pronounced dependence on W implies that the present approach provides a plausible range rather than a definitive “true” persistence value. Nevertheless, it is suitable for making comparisons across experimental contrasts.

Of course, our measure has some limitations. First, it is binary and does not account for extent of daily use, as demonstrated by the fact that patients who order batteries every 6 months are treated the same as patients who order every 18 months. More generally, the assumption of continued HA use over a period W after battery order is a simplification of the actual patient behavior that was introduced to derive a tractable measure. Second, it is a long-term measure that, in our dataset, cannot be used to investigate the change in HA use during the initial 18 months after HA fitting. Furthermore, since battery orders in the VHA are infrequent and irregular, they are insensitive to some behaviors (e.g., 39% of users classified as nonpersistent at $T=2$ years ordered batteries at a later date). Batteries obtained from sources outside VHA would lead to an underestimation of persistence. However, 99.8% of patients had at least one recorded battery order (Saunders et al. 2020) which implies that almost all patients order batteries through VHA. Persistence would also be underestimated if patients obtained a cochlear implant. However, few patients have procedural codes related to CIs (<0.2%), thus we consider this bias to be minor. Finally, rechargeable HAs are increasingly common, so battery-order data will become less

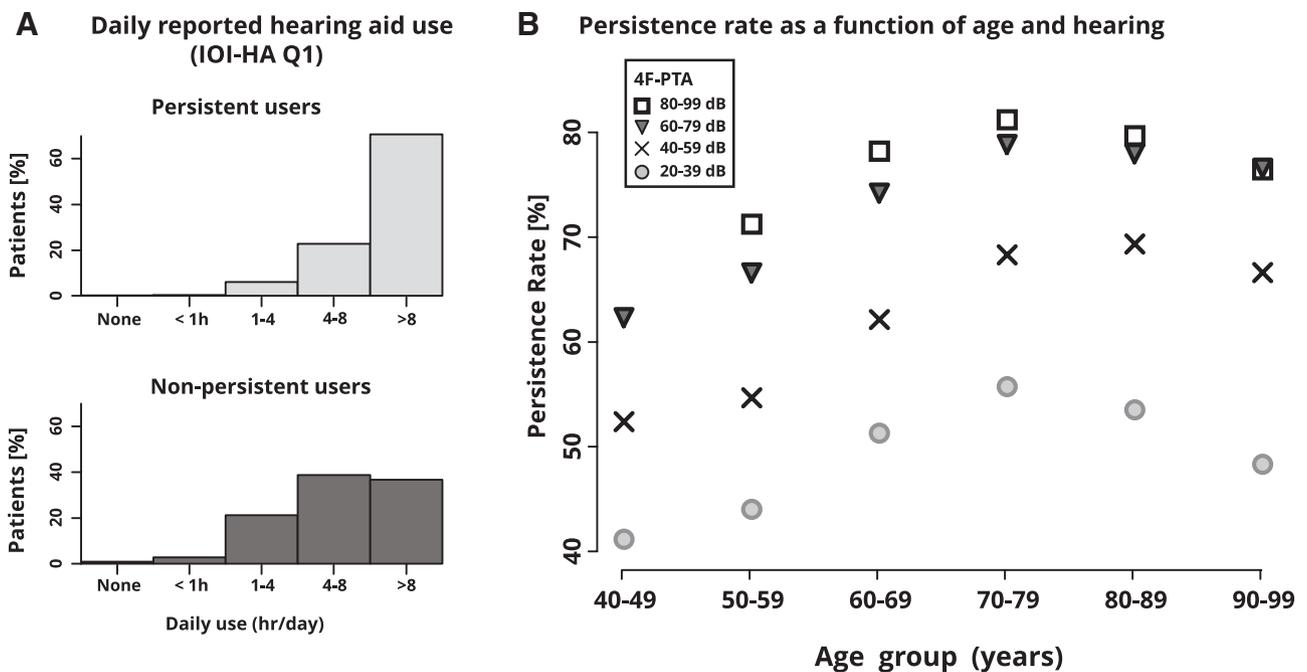


Fig. 3. Relationship between HA persistence and short-term HA usage as well as patient age and hearing. A, Self-reported short-term HA usage (IOI-HA question 1) for patients persistent and nonpersistent, respectively, at 2 years after HA fitting. B, Persistence rate as function of patient age and PTA. Data point for age group = 40–49 years; 4F-PTA = 80–99 dB omitted due to insufficient sample size. 4F-PTA, four-frequency pure-tone average; HA, hearing aid.

available as time goes on. In this case, our data suggest that a count of postfitting HA appointments might be an acceptable substitute. Furthermore, cloud-based data logging is becoming widespread, and might be a new way to track usage accurately.

A persistence measure could have been constructed in many ways. Key reasons for our choice were its simplicity and its similarity to measures frequently used in pharmacological research. For comparisons of the absolute rate of HA persistence found by our method against rates of medication persistence in other conditions, see Saunders et al. (2020).

In summary, we have presented a measure of long-term HA use persistence that forms a cornerstone of the analysis of our VHA dataset, and have provided evidence of its validity. We hope that this report will stimulate the use of this or similar research tools by others. By providing a deeper understanding of which factors affect HA persistence, such research might eventually improve clinical practice and service.

ACKNOWLEDGMENTS

The authors thank Kevin Quitmeyer, ShienPei Silverman, Kelly Reavis, Erin Robling, and M. Patrick Feeney for their support throughout the study.

This work was supported by the Medical Research Council (grant numbers MC_UU_00010/4, MR/S003576/1); and by the Chief Scientist Office of the

Scottish Government; VA Rehabilitation Research and Development grant #9230C, and by the VA Office of Academic Affiliations.

The authors have no conflicts of interest to disclose.

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Received November 4, 2020; accepted January 15, 2021

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