

1 Title

2 **First-line ovulation induction for polycystic ovary syndrome: an individual participant**
3 **data meta-analysis**

4 Running title

5 Ovulation induction for PCOS

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92 **Abstract**

93 **Background:** Polycystic ovary syndrome (PCOS) is the most frequent cause of anovulatory
94 infertility. In women with PCOS, effective ovulation induction serves as an important first-line
95 treatment for anovulatory infertility. Individual participant data (IPD) meta-analysis is
96 considered as the gold standard for evidence synthesis which provides accurate assessments of
97 outcomes from primary randomised controlled trials (RCTs) and allows additional analyses for
98 time-to-event outcomes. It also facilitates treatment-covariate interaction analyses and
99 therefore offers an opportunity for personalised medicine.

100 **Objective and rationale:** We aimed to evaluate the effectiveness of different ovulation
101 induction agents, in particular letrozole alone and clomiphene citrate (CC) plus metformin, as
102 compared to CC alone, as the first-line choice for ovulation induction in women with PCOS
103 and infertility, and to explore interactions between treatment- and participant-level baseline
104 characteristics.

105 **Search methods:** We searched electronic databases including MEDLINE, EMBASE and
106 Cochrane Central Register of Controlled Trials up to 20th December 2018. We included RCTs
107 comparing the following interventions with each other or placebo/ no treatment in women with
108 PCOS and infertility: CC, metformin, CC plus metformin, letrozole, gonadotrophin and
109 tamoxifen. We excluded studies on treatment-resistant women. The primary outcome was live
110 birth. We contacted the investigators of eligible RCTs to share the IPD and performed IPD
111 meta-analyses. We assessed the risk of bias by using the Cochrane risk of bias tool for RCTs.

112 **Outcomes:** IPD of 20 RCTs including 3962 women with PCOS were obtained. Six RCTs
113 compared letrozole and CC in 1284 women. Compared with CC, letrozole improved live birth
114 rates (3 RCTs, 1043 women, risk ratio [RR] 1.43, 95% confidence interval [CI] 1.17-1.75,
115 moderate-certainty evidence) and clinical pregnancy rates (6 RCTs, 1284 women, RR 1.45,
116 95% CI 1.23-1.70, moderate-certainty evidence), and reduced time-to-pregnancy (6 RCTs,

117 1235 women, hazard ratio [HR] 1.72, 95%CI 1.38-2.15, moderate-certainty evidence). Meta-
118 analyses of effect modifications showed a positive interaction between baseline serum total
119 testosterone levels and treatment effects on live birth (interaction RR 1.29, 95%CI 1.01-1.65).
120 Eight RCTs compared CC plus metformin to CC alone in 1039 women. Compared with CC
121 alone, CC plus metformin might improve clinical pregnancy rates (8 RCTs, 1039 women, RR
122 1.18, 95% CI 1.00-1.39, low-certainty evidence) and might reduce time-to-pregnancy (7 RCTs,
123 898 women, HR 1.25, 95%CI 1.00-1.57, low-certainty evidence), but there was insufficient
124 evidence of a difference on live birth rates (5 RCTs, 907 women, RR 1.08, 95% CI 0.87-1.35,
125 low-certainty evidence). Meta-analyses of effect modifications showed a positive interaction
126 between baseline insulin levels and treatment effects on live birth in the comparison between
127 CC plus metformin and CC (interaction RR 1.03, 95% CI 1.01-1.06).

128 **Wider implications:** In women with PCOS, letrozole improves live birth and clinical
129 pregnancy rates and reduces time-to-pregnancy compared to CC and therefore can be
130 recommended as the preferred first-line treatment for women with PCOS and infertility. CC
131 plus metformin may increase clinical pregnancy and may reduce time-to-pregnancy compared
132 to CC alone, while there is insufficient evidence of a difference on live birth. Treatments effects
133 of letrozole are influenced by baseline serum levels of total testosterone, while those of CC
134 plus metformin are affected by baseline serum levels of insulin. These interactions between
135 treatments and biomarkers on hyperandrogenaemia and insulin resistance provide further
136 insights into a personalised approach for the management of anovulatory infertility related to
137 PCOS.

138 **Key words:**

139 polycystic ovary syndrome, infertility, anovulation, ovulation induction, letrozole, clomiphene,
140 metformin, individual participant data, meta-analysis.

141

142 **Introduction**

143 Polycystic ovary syndrome (PCOS) is the most common endocrine disorder of reproductive
144 age women, and the prevalence among different geographic regions ranges from 5% to 21%,
145 depending on the criteria used (Lizneva, et al., 2016). PCOS is a heterogeneous syndrome
146 comprising of at least two of the following clinical characteristics according to the Rotterdam
147 diagnostic criteria: oligo-/ anovulation, clinical and/or biochemical hyperandrogenism, or
148 polycystic ovaries morphology based on ultrasound assessment (Rotterdam ESHRE/ASRM-
149 Sponsored PCOS Consensus Workshop Group, 2004).

150 Anovulatory infertility is usually one of the key features that women with PCOS are confronted
151 with. Simple and effective infertility treatments as the first-line choice are therefore important.
152 Our previous network meta-analysis compared available first-line treatment options for women
153 with PCOS with infertility and found that letrozole and combined clomiphene citrate (CC)-
154 metformin were superior to other ovulation induction medications in terms of clinical
155 pregnancy and that letrozole resulted in more live births than other interventions, including CC
156 (Wang, et al., 2017). These findings are in agreement with the evidence summarised in the
157 International evidence based guideline for the assessment and management of PCOS (Teede,
158 et al., 2018).

159 As women with PCOS represent a heterogeneous population according to the diagnostic
160 criteria, it is important to identify which individuals benefit most from a particular treatment
161 so that clinicians can provide personalised care (Wang and Mol, 2017). However, primary
162 RCTs are usually underpowered to detect subgroup effects (Riley, et al., 2010). Subgroup
163 analyses in meta-analyses of aggregate data are at risk of ecological bias due to the ignorance
164 of within-study interactions, or are even impossible to perform due to heterogeneous reporting
165 of subgroup data in the primary trials (Riley, et al., 2010).

166 Moreover, time-to-pregnancy is also an important patient-centred outcome, but it has never
167 been reported in previous meta-analyses on PCOS. This is likely due to the unavailability of
168 the data in the publication as well as the methodological challenges on data extraction and
169 synthesis. In addition, the primary trials are not always of high quality in terms of analyses and
170 reports (Eshre Capri Workshop Group, 2018), which can directly affect the data extraction,
171 analysis and risk of bias assessment process in subsequent meta-analyses.

172 These deficiencies in aggregate data meta-analyses can potentially be overcome by using
173 individual participant data (IPD). IPD meta-analysis has been described as the gold standard in
174 evidence synthesis, by engaging investigators of the primary trials to provide the raw data of
175 the primary trials (Broeze, et al., 2010). Such strategy facilitates derivation of the information
176 beyond the primary publication, standardisation of inclusion criteria, outcomes and analyses
177 across trials, and investigations of subgroup effects and time-to-event outcomes. (Broeze, et
178 al., 2010, Riley, et al., 2010).

179 We therefore performed an IPD meta-analysis to evaluate the effectiveness of different
180 ovulation induction agents, in particular letrozole alone and CC plus metformin, as compared
181 to CC alone, as the first-line choice for ovulation induction in women with PCOS and
182 infertility, and to explore interactions between treatment- and participant-level baseline
183 characteristics.

184

185 **Methods**

186

187 **Registration and literature search**

188 This IPD meta-analysis was conducted based on a registered protocol (PROSPERO
189 CRD42017059251) and reported according to the Preferred Reporting Items for Systematic

190 Review and Meta-Analyses of individual participant data (PRISMA-IPD) statement (Stewart,
191 et al., 2015).

192 We updated the searches in MEDLINE, EMBASE and Cochrane Central Register of
193 Controlled Trials in September 2017, based on our previous search strategies for a network
194 meta-analysis on treatment strategies for World Health Organization (WHO) II anovulation
195 (Wang, et al., 2017). In brief, the search terms included both index terms as well as free words
196 on PCOS, anovulation and ovulation induction. After completing data requesting process, we
197 further updated the search on 20th December 2018 to identify the latest studies. We also
198 searched the WHO International Clinical Trials Registry Platform (WHO ICTRP) and U.S.
199 National Institutes of Health (clinicaltrials.gov) and ISRCTN registry to identify ongoing trials.
200 In addition, we reviewed the references lists of relevant papers and corresponded with trialists
201 in PCOS to identify potential eligible trials that we might have missed.

202

203 **Eligibility criteria**

204 We included RCTs comparing the following interventions with each other or placebo/no
205 treatment: clomiphene citrate (CC), metformin, CC and metformin combined, letrozole,
206 gonadotrophins and tamoxifen in women with WHO II anovulation, including PCOS. We
207 excluded trials reporting on treatment-resistant women, trials comparing different doses of the
208 same intervention and quasi-RCTs. We did not apply language restrictions. For crossover trials,
209 we only included the data in the first phase.

210 The primary outcome was live birth. The secondary outcomes were clinical pregnancy,
211 ovulation, miscarriage, multiple pregnancy and time to pregnancy.

212

213 **Study selection and data collection**

214 Two members of the review team (from RW, WL and EMB) independently assessed the titles
215 and abstracts to exclude irrelevant studies and subsequently reviewed the full-text articles to
216 evaluate their eligibility. Disagreements were resolved by discussion with a third author
217 (BWM, MvW or RJN).

218 We contacted investigators of eligible RCTs to share the de-identified IPD and established the
219 International Ovulation Induction IPDMA Collaboration. We sent at least two more reminders
220 when we did not receive responses.

221 We obtained de-identified IPD including baseline characteristics including age, body mass
222 index (BMI), ethnicity, type of infertility (primary/secondary), treatment history (treatment-
223 naïve or not), fasting glucose, fasting insulin, total testosterone, sex hormone binding globulin
224 (SHBG), ovarian volume and the Ferriman-Gallwey score for hirsutism. We also obtained data
225 on allocated treatments, number of ovulation induction cycles, ovulation and fertility outcomes
226 including live birth, clinical pregnancy, miscarriage and multiple pregnancy.

227 We checked data for consistency by comparing the analyses from obtained IPD with the
228 original publications. We discussed any inconsistencies or obvious errors with investigators of
229 primary RCTs and solved discrepancies by consensus.

230

231 **Risk of bias assessment**

232 Two members of the review team independently evaluated the risk of bias in each included
233 RCT, using the domain-based evaluation tool described in the Cochrane Handbook for
234 Systematic Reviews of Interventions (Higgins and Green, 2011). We assessed the following
235 domains as low risk of bias, unclear or high risk of bias: random sequence generation,
236 allocation concealment, blinding of participants and personnel, blinding of outcome assessors,
237 incomplete outcome data, selective reporting (reporting bias) and other sources of bias. When

238 the risk of bias for a domain was unclear, investigators of these RCTs were asked to provide
239 additional information to resolve the uncertainty.

240 We assessed the overall certainty of the evidence across RCTs by using the Grading of
241 Recommendations Assessment, Development and Evaluation (GRADE) approach, including
242 the risk of bias, consistency of effect, imprecision, indirectness and publication bias.

243

244 **Data synthesis**

245 We conducted all analyses based on an intention-to-treat principle using woman randomised
246 per allocated group as the unit of all analyses. We performed two-stage random-effects IPD
247 meta-analyses for letrozole versus CC alone and CC with metformin versus CC alone. For
248 dichotomous outcomes, we calculated risk ratios (RRs) and 95% confidence intervals (CIs) and
249 presented statistical heterogeneity by using I^2 statistic (Higgins and Green, 2011). For time-to-
250 event outcomes, we used the number of treatment cycles as an approximate estimate for time
251 and visualised the summary time-to-event in simple non-stratified Kaplan-Meier curves. We
252 also estimated hazard ratios (HR) in Cox proportional hazards regression models for discrete
253 time and pooled HRs and 95% CI, by using the generic inverse variance method (Fisher, 2015).
254 Subgroup effects were estimated for the primary outcome by treatment-covariate interaction
255 terms within trials and subsequent meta-analyses of interactions, as interactions using within-
256 trials information alone without considering between-trials interactions are recommended as
257 the standard practice to avoid ecological bias (Fisher, et al., 2017). We explored the treatment-
258 covariate interactions of the following pre-specified baseline covariates: age, BMI, ethnicity,
259 primary/secondary infertility, treatment history, hirsutism score, insulin resistance (serum
260 glucose and insulin level), hyperandrogenaemia status (testosterone, SHBG, free androgen
261 index) and ovarian volume. We also added the analysis of homeostatic model assessment for
262 insulin resistance (HOMA-IR) as requested during the peer review process. For dichotomous

263 covariates with statistically significant interaction, we further performed stratified analyses to
264 illustrate the treatment effects in different strata of the subgroups. Continuous variables were
265 analysed as such without categorisation. For continuous covariates with statistically significant
266 interaction, we further presented a weighted mean curve and pointwise confidence interval
267 based on treatment-covariate interactions estimated in relevant studies. Due to the potential
268 type I error, the results of subgroup analyses were all considered exploratory.

269 To evaluate the IPD availability bias, we performed a network meta-analysis of RCTs with IPD
270 in a random-effects multivariate meta-analysis model (Riley, et al., 2017, White, 2015) on live
271 birth and clinical pregnancy, and then compared the results with a network meta-analysis of all
272 eligible RCTs. If these results were consistent, we considered the included RCTs with IPD
273 representative of all the eligible RCTs.

274 We performed a sensitivity analysis on studies with low risk of bias in allocation concealment
275 as planned. As the majority of eligible studies focused only on treatment-naïve women with
276 PCOS, these studies did not contribute to within-study interaction for treatment history and
277 were not included in the treatment-covariate analysis. We performed a post-hoc sensitivity
278 analysis by including only treatment-naïve women to demonstrate the robustness of the results.
279 We conducted all the analyses in Stata software version 15.1 (Stata Corp, College Station, TX,
280 USA).

281

282 **Results**

283

284 **Characteristics of included studies**

285 The final updated search yielded 709 non-duplicated studies (Figure 1). After screening the
286 titles and abstracts, 636 irrelevant studies were excluded. Finally, a total of 62 studies (61
287 publications, 9356 women) fulfilled the inclusion criteria and were included. These studies

288 were published in English (n=58), French (n=1) (Boudhraa, et al., 2010), Italian (n=1)
289 (Santonocito, et al., 2009), Turkish (n=1) (Aygen, et al., 2007) and Persian (n=1) (Lorzadeh,
290 et al., 2011).

291 IPD was not sought from eight studies (575 women), due to insufficient contact information
292 (n=6; 359 women) (Beigi, 2006, Boudhraa, et al., 2010, Cudmore and Tupper, 1966, El-Biely
293 and Habba, 2001, Garcia, et al., 1985, Johnson, et al., 1966) or because the studies were
294 identified after our data requesting timeline (n=2; 216 women) (Fatima, et al., 2018, Topçu, et
295 al., 2017). For the remaining 54 studies (8781 women), the primary investigators were
296 contacted to share IPD of the primary studies. IPD from 34 studies (4819 women) were not
297 available, due to no response (n=23; 3258 women) (Abuelghar, et al., 2013, Atay, et al., 2006,
298 Ayaz, et al., 2013, Banerjee Ray, et al., 2012, Basirat, et al., 2012, Boostanfar, et al., 2001,
299 Chen, et al., 2016, Dasari and Pranahita, 2009, Dehbashi, et al., 2009, Hossein-Rashidi, et al.,
300 2016, Jahan, 2015, Karimzadeh, et al., 2007, Karimzadeh and Javedani, 2010, Lopez, et al.,
301 2004, Lorzadeh, et al., 2011, Maged, et al., 2015, Robinson, et al., 2003, Roy, et al., 2012,
302 Selim and Borg, 2012, Seyedoshohadaei, et al., 2012, Sharief and Nafee, 2015, Sheikh-El-Arab
303 Elsedek and Elmaghraby, 2011, Zeinalzadeh, et al., 2010), data loss (n=10; 1411 women)
304 (Aygen, et al., 2007, Badawy, et al., 2009, Badawy and Gibreal, 2011, Fleming, et al., 2002,
305 Keikha and Shahraki Mojahed, 2011, Khorram, et al., 2006, Mobusher, 2014, Santonocito, et
306 al., 2009, Tang, et al., 2006, Zain, et al., 2009) or legal reasons (n=1; 150 women) (Moussa, et
307 al., 2016). The details of these studies are listed in Supplementary Table 1.

308 IPD were available for at least one outcome from 20 studies (3962 women), including three
309 from the US (Legro, et al., 2007, Legro, et al., 2014, Williams, et al., 2009), three from Italy
310 (Leanza, et al., 2014, Palomba, et al., 2005, Vegetti, et al., 1999), three from Turkey (Bayar, et
311 al., 2006, Nazik and Kumtepe, 2012, Sahin, et al., 2004), two from the UK (Amer, et al., 2017,
312 Lord, et al., 2006), two from China (Liu, et al., 2017, Wu, et al., 2017), two from India (Kar,

313 2012, Kar and Sanchita, 2015), two studies (in one publication) from New Zealand (Johnson,
314 et al., 2010), one from The Netherlands (Moll, et al., 2006), one from Finland (Morin-Papunen,
315 et al., 2012) and one from multiple countries (The Netherlands, UK, Malta, Belgium, Argentina
316 and Colombia) (Homburg, et al., 2012). These RCTs were published in English between 1999
317 and 2017, with 11 (55%) published after 2010.

318 Participants in all 20 RCTs were women with PCOS. In one RCT, participants were diagnosed
319 with PCOS by fulfilling at least three of the following: PCO morphology, oligo/amenorrhoea,
320 hirsutism, hyperandrogenaemia and elevated serum LH/FSH ratio (Sahin, et al., 2004); while
321 in the remaining 19 RCTs, the participants were women with PCOS based on the Rotterdam
322 criteria (Bayar, et al., 2006, Kar, 2012, Leanza, et al., 2014, Liu, et al., 2017, Nazik and
323 Kumtepe, 2012) or different phenotypes, including Phenotype B (ovulatory dysfunction +
324 androgen excess) (Amer, et al., 2017, Homburg, et al., 2012, Johnson, et al., 2010, Kar and
325 Sanchita, 2015, Legro, et al., 2007, Legro, et al., 2014, Lord, et al., 2006, Morin-Papunen, et
326 al., 2012, Palomba, et al., 2005, Williams, et al., 2009, Wu, et al., 2017) or Phenotype D
327 (ovulatory dysfunction + PCO) (Moll, et al., 2006, Vegetti, et al., 1999).

328 For RCTs involving two stages of different interventions, including cross-over studies, we only
329 included the data in the first stage. We included the IPD comparing letrozole versus CC before
330 crossing over (Amer, et al., 2017) and included the IPD comparing metformin versus placebo
331 within the first three months before starting other ovulation induction agents (Morin-Papunen,
332 et al., 2012). In one RCT (Nazik and Kumtepe, 2012), switching between intervention and the
333 control after the first cycle was allowed during the trial and the analysis in the primary
334 publication was on a per-cycle basis; and therefore we only included the IPD of the first cycle.
335 In summary, four RCTs compared three interventions (CC plus metformin or CC alone versus
336 metformin (Johnson, et al., 2010, Kar and Sanchita, 2015, Legro, et al., 2007) or CC with
337 metformin or letrozole versus CC (Liu, et al., 2017)) and the remaining 16 compared two

338 interventions. The most common comparisons were CC with metformin versus CC alone (8
339 RCTs) (Johnson, et al., 2010, Kar and Sanchita, 2015, Leanza, et al., 2014, Legro, et al., 2007,
340 Liu, et al., 2017, Moll, et al., 2006, Sahin, et al., 2004, Williams, et al., 2009) and letrozole
341 versus CC alone (6 RCTs) (Amer, et al., 2017, Bayar, et al., 2006, Kar, 2012, Legro, et al.,
342 2014, Liu, et al., 2017, Nazik and Kumtepe, 2012).

343

344 **Quality of evidence of individual studies**

345 The details of risks of bias assessments within individual studies are presented in Figure 2. All
346 RCTs (n=20) reported adequate methods of random sequence generation. Sixteen RCTs (80%)
347 reported adequate methods of allocation concealment while the other four used an open
348 allocation schedule without concealment (Kar, 2012, Kar and Sanchita, 2015, Liu, et al., 2017,
349 Nazik and Kumtepe, 2012). Fourteen RCTs (70%) blinded the participants and personnel
350 during the trial while six RCTs applied an open label design (Homburg, et al., 2012, Kar, 2012,
351 Kar and Sanchita, 2015, Liu, et al., 2017, Nazik and Kumtepe, 2012, Vegetti, et al., 1999).
352 Given that all outcomes of interest were objective outcomes, it is unlikely that the non-blinded
353 design will affect the outcome measurement and therefore detection bias was rated at low risk
354 for all the included studies. One RCT (5%) had high risk of attrition bias, with 22% overall
355 missing outcome data and 31% missing outcome data in the metformin group (Kar and
356 Sanchita, 2015). One RCT (5%) was at another risk of bias due to allowing imbalanced co-
357 intervention (CC) in both groups.

358

359 **Meta-analyses of letrozole versus CC**

360 ***Live birth***

361 IPD were available in six RCTs comparing letrozole and CC, including 1284 women with
362 PCOS. The forest plot of IPD Meta-analysis on live birth is presented in Figure 3a. Compared

363 with CC, letrozole increased live birth rates (3 RCTs, 1043 women, RR 1.43, 95% CI 1.17-
364 1.75, $I^2=0$, moderate certainty of evidence). Sensitivity analysis on studies with low risk of bias
365 at allocation concealment and on treatment-naïve women were consistent with the main
366 findings (2 RCTs, 909 women, RR 1.42, 95% CI 1.14-1.76, $I^2=0$; 3 RCTs, 627 women, RR
367 1.41, 95%CI 1.11-1.79, $I^2=0$) (Supplementary Table 2).

368 ***Secondary outcomes***

369 Compared with CC alone, letrozole improved clinical pregnancy (6 RCTs, 1284 women, RR
370 1.45, 95%CI 1.23-1.70, $I^2=0$, moderate certainty of evidence, Figure 3b) and ovulation rates (5
371 RCTs, 1210 women, RR 1.13, 95%CI 1.07-1.20, $I^2=0$, moderate certainty of evidence, Table
372 2). There was insufficient evidence of a difference between letrozole and CC alone in terms of
373 multiple pregnancy or miscarriage (Table 2).

374 The summary Kaplan-Meier curve for time to pregnancy is presented in Figure 4a. Subsequent
375 pooled analysis of HRs showed that compared to CC, letrozole improved time-to-pregnancy (6
376 RCTs, 1235 women, HR 1.72, 95%CI 1.38-2.15, $I^2=0$, moderate certainty of evidence).

377 ***Treatment-covariate interactions***

378 A meta-analysis of effect modifications showed a positive interaction between baseline serum
379 total testosterone levels and treatment effects on live birth in the comparison between letrozole
380 and CC (interaction RR 1.29, 95%CI 1.01-1.65, 3 RCTs, 1039 women, Figure 5a). This
381 suggests that women with a higher baseline serum total testosterone level have a larger
382 treatment effect of letrozole versus CC on live birth, compared to women with a lower baseline
383 serum total testosterone level. Such an interaction was consistent across studies ($I^2=0$). To
384 directly illustrate the association between baseline serum total testosterone level and relative
385 treatment effects, this interaction is also presented in a weighted mean curve with 95% CI
386 (Figure 5b). Meta-analysis did not find any other treatment-covariate interactions (Table 3).

387

388 **Meta-analyses of CC plus metformin versus CC**

389 ***Live birth***

390 IPD were available in eight RCTs comparing CC with metformin and CC alone, including 1039
391 women with PCOS. The forest plot of IPD Meta-analysis on live birth is presented in Figure
392 3c. Compared with CC alone, there was insufficient evidence of a difference between CC with
393 metformin and CC alone on live birth (5 RCTs, 907 women, RR 1.08, 95%CI 0.87-1.35,
394 $I^2=5.6%$, low certainty of evidence). Sensitivity analyses on studies with low risk of bias at
395 allocation concealment and on treatment-naïve women showed very small treatment effects
396 with wide CIs (3 RCTs, 714 women, RR 1.02, 95%CI 0.76-1.37, $I^2=33.2%$; 5 RCTs, 662
397 women, RR 1.06, 95%CI 0.83-1.34, $I^2=3.9%$) (Supplementary Table 2).

398 ***Secondary outcomes***

399 Compared with CC alone, CC with metformin might improve clinical pregnancy (8 RCTs,
400 1039 women, RR 1.18, 95% CI 1.00-1.39, $I^2=6.9%$, low certainty of evidence, Figure 3b).
401 There was insufficient evidence of a difference between CC with metformin and CC alone on
402 ovulation, multiple pregnancy or miscarriage (Table 2).

403 The summary Kaplan-Meier curve is presented in Figure 4b. Pooled analysis of HRs showed
404 that compared to CC alone, CC with metformin might improve time-to-pregnancy (7 RCTs,
405 898 women, HR 1.25, 95%CI 1.00-1.57, $I^2=0$, low certainty of evidence).

406 ***Treatment-covariate interactions***

407 Meta-analyses of effect modifications showed a positive interaction between baseline insulin
408 levels and treatment effects on live birth in the comparison between CC with metformin and
409 CC alone (interaction RR 1.03, 95%CI 1.01-1.06, 4 RCTs, 741 women, Figure 5c). Such an
410 interaction was consistent across studies ($I^2=0$). This suggests that women with a higher
411 baseline serum insulin level have larger treatment effects of CC with metformin versus CC
412 alone on live birth, compared to women with a lower baseline serum insulin level. Such an

413 interaction was also presented in a weighted mean curve with 95%CI (Figure 5d). Additional
414 meta-analysis of interactions for HOMA-IR was performed as requested during the peer review
415 process and it also showed a positive interaction between baseline HOMA-IR and treatment
416 effects on live birth in the comparison between CC with metformin and CC alone (interaction
417 RR 1.14, 95%CI 1.03-1.25, 4 RCTs, 736 women, $I^2=0$, Table 3). Meta-analyses did not find
418 any other treatment-covariate interactions (Table 3).

419

420 **IPD availability bias**

421 With regards to IPD availability bias, network meta-analyses of 20 RCTs with IPD showed
422 similar results to network meta-analyses of all eligible RCTs on both live birth and clinical
423 pregnancy (Supplementary Table 3). Therefore, the participants in RCTs with IPD were
424 representative of all the eligible participants with PCOS. The transitivity assumption of
425 network meta-analyses was considered valid as the interventions of interest and placebo/no
426 treatment were jointly randomisable.

427

428 **Discussion**

429

430 **Summary of evidence**

431 This IPD meta-analysis showed that in women with PCOS, letrozole increased live birth rates
432 compared to CC alone and the overall certainty of evidence was moderate. Such treatment
433 benefits of letrozole compared to CC alone were more predominant in women with higher
434 baseline serum levels of total testosterone. There was insufficient evidence of a difference
435 between CC plus metformin and CC alone in live birth rates and the overall certainty of
436 evidence was low, mainly due to risk of bias and imprecision. The potential benefit of CC in
437 combination with metformin compared to CC alone were more pronounced in women with

438 higher baseline serum insulin or HOMA-IR levels. We did not find other treatment-covariate
439 interactions on live birth for other prespecified covariates including age, BMI, ethnicity,
440 primary/secondary infertility, treatment history, Ferriman–Gallwey score for hirsutism, SHBG,
441 free androgen index, fasting glucose levels or ovarian volume.

442

443 **Strengths and limitations**

444 Establishing the International Ovulation Induction IPDMA Collaboration facilitated a platform
445 for key trialists in PCOS to collaborate and share the IPD of the primary trials. It provided us
446 the opportunity to collect unpublished information of the primary trials including the details of
447 randomisation and allocation concealment, treatment history, subgroup data and time-to-
448 pregnancy. Such information allowed us to assess the quality of included trials precisely, to
449 investigate treatment-covariate interactions and to take account of the time in the analyses. The
450 findings of this IPD meta-analysis provide the best available up-to-date evidence.

451 Moreover, we applied a comprehensive search strategy without language restrictions and
452 updated the search after completing data requesting in case we missed the most recent RCTs.
453 Of the newly identified RCTs, one compared CC plus metformin vs CC in 128 women but did
454 not report live birth (Fatima, et al., 2018), while the other one compared tamoxifen vs CC in
455 88 women (Topçu, et al., 2017). Although we did not seek IPD from two RCTs identified after
456 the data requesting deadline, adding IPD of these two studies is unlikely to change the main
457 findings.

458 In addition, the investigation of subgroup effects includes within-study interaction only
459 according to current statistical practice for IPD meta-analyses (Fisher, et al., 2017) and
460 therefore are free from ecological bias. For continuous covariates, without categorisation of the
461 data, the statistical power was not compromised. Further illustration of interactions in weighted
462 mean curve makes the interactions easier to interpret.

463 Nevertheless, this IPD meta-analysis has a few limitations. First, we were not able to access
464 the IPD of all eligible studies. IPD were available for 32% (20/62) of the included trials,
465 comprising 42% (3962/9356) of the eligible women with PCOS and the proportions of IPD
466 availability was higher for studies reporting live birth (44% trials including 65% eligible
467 women, Supplementary Table 3). This seems to be partly due to the long history of research
468 on ovulation induction, with the first trial published in 1966. We were however able to access
469 IPD of the highest-quality trials published within the last 15 years and we did not detect
470 evidence of availability bias. Second, most of the planned subgroup analyses were based on
471 two to three of the included studies and therefore may still be underpowered due to the
472 unavailability of data on relevant covariates and/or live birth. Some primary trials only included
473 a relatively homogeneous ethnicity group and therefore IPD in such trials could not contribute
474 to the analysis of treatment-ethnicity interaction as no within-trial interaction was available.
475 Third, as treatment-resistant women were excluded from this IPD meta-analysis, the findings
476 can be applied in clinical practice on the choice of first-line treatment only. Last, we planned a
477 one-stage IPD meta-analysis in the protocol but decided to use a two-stage approach before the
478 final analysis. A two-stage approach allows graphical presentations for both overall treatment
479 effects and treatment-covariate interactions, which is important for clinical interpretation,
480 while it is not obvious how best to present graphically the results of a one-stage model (Fisher,
481 et al., 2017). In addition, the two-stage approach automatically avoids ecological bias by
482 accounting for within-trial interactions only (Fisher, et al., 2017). Given the relatively large
483 number of participants, low heterogeneity and overall good to moderate quality of included
484 studies, we would expect both approaches to give very similar results.

485

486 **Interpretations and clinical implications**

487 The overall effects of letrozole and CC plus metformin vs CC on live birth and clinical
488 pregnancy in this IPD meta-analysis were in agreement with existing systematic reviews
489 (Franik, et al., 2018, Morley, et al., 2017, Wang, et al., 2017) as well as the most recent the
490 international evidence-based guideline recommendations (Teede, et al., 2018). Based on the
491 findings of this IPD meta-analysis, letrozole can be recommended as the first-line ovulation
492 induction medication in women with PCOS and infertility, provided off-label use is allowed
493 and women are fully informed. Compared to CC alone, CC plus metformin may increase
494 clinical pregnancy rates but the evidence on live birth was insufficient. Sensitivity analysis
495 showed that the treatment effects on live birth seemed very small. The discrepancies between
496 clinical pregnancy and live birth were likely due to the bias arising from low quality of studies
497 which did not report live birth. Further evidence is needed to address this question.

498 Subgroup analyses showed that women with higher baseline serum levels of total testosterone
499 may benefit more from letrozole compared to CC and women with higher baseline serum levels
500 of insulin may benefit more from CC plus metformin compared to CC alone. Such positive
501 interactions were consistent across trials and supported from a biological perspective. Letrozole
502 has been introduced as an ovulation induction agent since 2001 and it inhibits aromatase,
503 therefore increasing gonadotropin secretion by release of the hypothalamic/pituitary axis from
504 estrogenic negative feedback and resulting in stimulation of ovarian follicle development
505 (Mitwally and Casper, 2001). According to the recent “two triangles hypothesis” for
506 folliculogenesis in PCOS, pre-antral follicle growth is excessive due to intrinsic androgen
507 excess that renders granulosa cells hypersensitive to FSH, with consequently excessive AMH
508 expression (Dewailly, et al., 2016) Therefore, hyperandrogenaemia may improve the response
509 to letrozole by enhancing the sensitivity of FSH receptors. However, such an interaction was
510 not observed in other biomarkers of hyperandrogenaemia or hirsutism. This is likely due to the
511 fact that the severity of hirsutism does not correlate well with the magnitude of androgen

512 excess, as hirsutism is an expression of hyperandrogenism on hair follicles mediated through
513 different pathways from those affecting the ovaries and follicles (Escobar-Morreale, et al.,
514 2012). Metformin is an insulin sensitising agent that decreases gluconeogenesis and lipogenesis
515 and enhances peripheral glucose uptake and therefore increases insulin sensitivity (Naderpoor,
516 et al., 2015). The addition of metformin may further improve insulin resistance in women with
517 higher fasting insulin or HOMA-IR levels and therefore improve pregnancy outcomes. We
518 acknowledge that insulin levels are affected by many factors, ranging from physical activity
519 and pre-test duration of fasting to sample handling and assay variability (Cassar, et al., 2016).
520 Therefore the international evidence-based guideline does not recommend clinical
521 measurement of insulin resistance at present due to the lack of accuracy (Teede, et al., 2018).
522 In addition, SHBG has been proposed as a measure of insulin resistance (Cassar, et al., 2016),
523 but the findings in our IPD meta-analysis did not support treatment-by-SHBG interactions. Our
524 work provides preliminary evidence that there may be a role for assessing insulin resistance in
525 PCOS and infertility and supports the need to assess insulin resistance in infertility studies.
526 We did not find ethnicity differences on treatment effects. This could be partly due to self-
527 reported ethnicity without objective or DNA validation in all trials. We also did not find other
528 treatment-covariate interactions on live birth for other prespecified covariates including age,
529 BMI, primary/secondary infertility, treatment history, Ferriman–Gallwey score for hirsutism,
530 SHBG, free androgen index, fasting glucose levels or ovarian volume. Although analyses of
531 subgroup effects were prespecified in the protocol, these results should still be considered
532 exploratory due to multiplicity.

533 Time is an important measurement for infertility outcomes, especially in the assessment of the
534 effectiveness of multi-cycle treatments. However, time-to-event outcomes have seldomly been
535 reported in meta-analyses of infertility trials as fertility outcomes are usually considered as
536 dichotomous outcomes and Kaplan-Meier curves are rarely presented. Our IPD meta-analysis

537 used number of cycles as a measure of time and evaluated time-to-pregnancy by estimating
538 HRs and presenting summary Kaplan-Meier curves. Time-to-event analysis takes time and
539 censored participants into account and provides more accurate estimates of treatment effect.
540 Our analyses on time-to-pregnancy were inconsistent with those of clinical pregnancy.

541

542 **Research implications**

543 IPD meta-analyses are useful to inform the design, conduct, analysis, and interpretation of trials
544 (Tierney, et al., 2015). Given the consistent treatment benefits of letrozole across different
545 fertility outcomes, future trials investigating new interventions for PCOS should choose
546 letrozole as the reference arm. New trials are encouraged to incorporate treatment selection
547 markers in their design to guide treatment decision (Janes, et al., 2011), and the impact of these,
548 including age, BMI and other biomarkers, need to be confirmed in future trials. More
549 specifically, biomarkers for hyperandrogenaemia and insulin resistance could be applied in
550 trials that evaluate metformin. Due to the limited accuracy for measuring existing insulin
551 resistance biomarkers, optimal methods to assess insulin resistance in future trials should also
552 be considered.

553 Developing and implementing a core outcome set for infertility (Duffy, et al., 2018) and PCOS
554 should be recommended to ensure outcomes are reported and collected consistently across
555 future trials on infertility and PCOS to reduce research waste .

556

557 **Conclusions**

558 Our IPD meta-analysis shows that in women with PCOS, letrozole improves live birth and
559 clinical pregnancy rates and reduces time-to-pregnancy compared to CC alone. CC plus
560 metformin may improve clinical pregnancy rates and may reduce time-to-pregnancy compared
561 to CC alone, but there is insufficient evidence of a difference on live birth.

562 Treatments effects of letrozole are influenced by baseline serum levels of total testosterone
563 while those of CC plus metformin are affected by baseline serum levels of insulin. These
564 interactions between treatments and biomarkers on hyperandrogenaemia and insulin resistance
565 provide further insights into a personalised approach towards the clinical management of
566 anovulatory infertility related to PCOS and therefore should be confirmed in future studies.
567

568 **Authors' roles**

569 RW, RSL, SB, RJN, MvW and BWM conceptualised and designed the study. RW, WL, EMB,
570 RJN, MvW and BWM collected the data. RSL, HZ, XW, JG, LMP, RH, TEK, EM, SK, WH,
571 NPJ, SAA, WV, SP, AF, UO, HN, CDW, GF, JL and YS provided and interpreted data from
572 the included trials. RW, WL, EMB, MvW and BWM cleaned and analysed the data. RW
573 drafted the first manuscript. All authors interpreted the pooled data, critically revised the
574 manuscript for important intellectual content, and approved the final version.

575

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598 **Conflict of interest**

599 RSL reports consultancy fees from Abbvie, Bayer, Fractyl and Ogeda and research sponsorship
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873

874 Figure legends

875 Figure 1. PRISMA-IPD flow diagram

876 Figure 2. Risk of bias assessments of individual RCTs

877 Figure 3. Meta-analyses of letrozole versus CC and CC plus metformin versus CC on live
878 birth and clinical pregnancy

879 Figure 4. Summary Kaplan-Meier curves for time-to-event outcomes

880 Figure 4a-4b illustrate the non-stratified summary Kaplan-Meier curves for time-to-
881 pregnancy in the comparisons of letrozole versus CC and CC plus metformin versus CC,
882 respectively.

883 Participants with pregnancy before the first treatment cycles were not included in the
884 'Numbers at risk' table below and data were not stratified by trial in this Kaplan-Meier curve.

885 The figures were intended to visualise time-to-event outcomes, but not to show statistical
886 significance.

887 Figure 5. Forest plots and weighted mean curves for treatment-covariate interactions

888 5a. Forest plot of interactions between baseline serum total testosterone (TT) level and effect
889 of letrozole versus CC on live birth.

890 5b. Weighted mean curve with pointwise 95% CI of interactions between baseline serum total
891 testosterone level and relative effect of letrozole versus CC on live birth. 5c. Forest plot of
892 interactions between baseline serum insulin level and effect of CC plus metformin versus CC
893 on live birth.

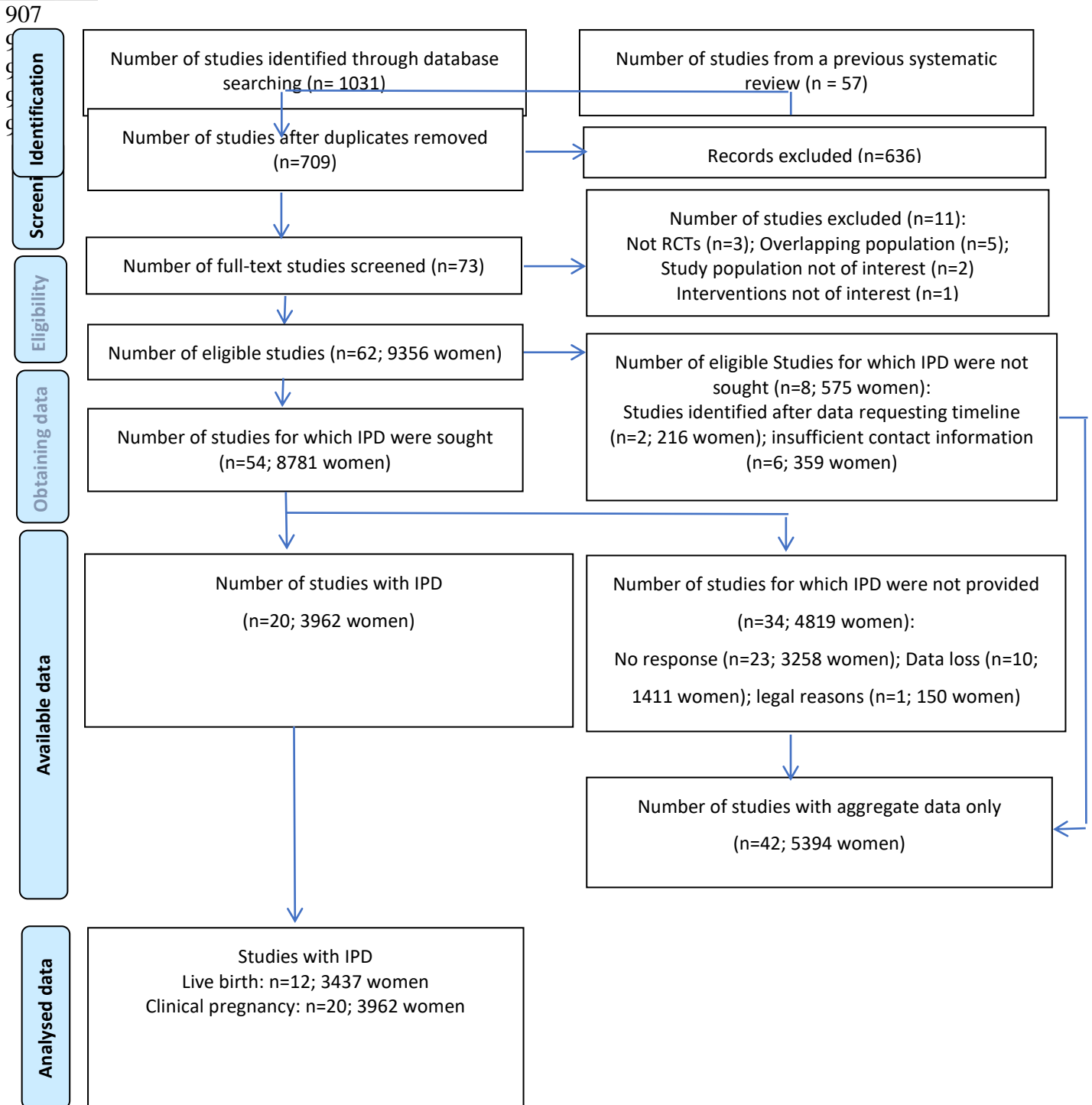
894 5d. Weighted mean curve with pointwise 95% CIs of interactions between baseline serum
895 insulin level and effect of CC plus metformin versus CC on live birth.

896 5a,c. Circles are used to depict the interaction effects within individual trials as well as the
897 overall interaction effect. The sizes of the circles are in proportion to the inverse of the
898 variance of the estimates.

899 5b,d. Blue line represents for the weighted mean effect of covariate on log risk ratios in the
900 comparison between letrozole and CC. Red lines represent for pointwise 95% CI of
901 interactions.
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PRISMA IPD Flow Diagram



The PRISMA IPD flow diagram

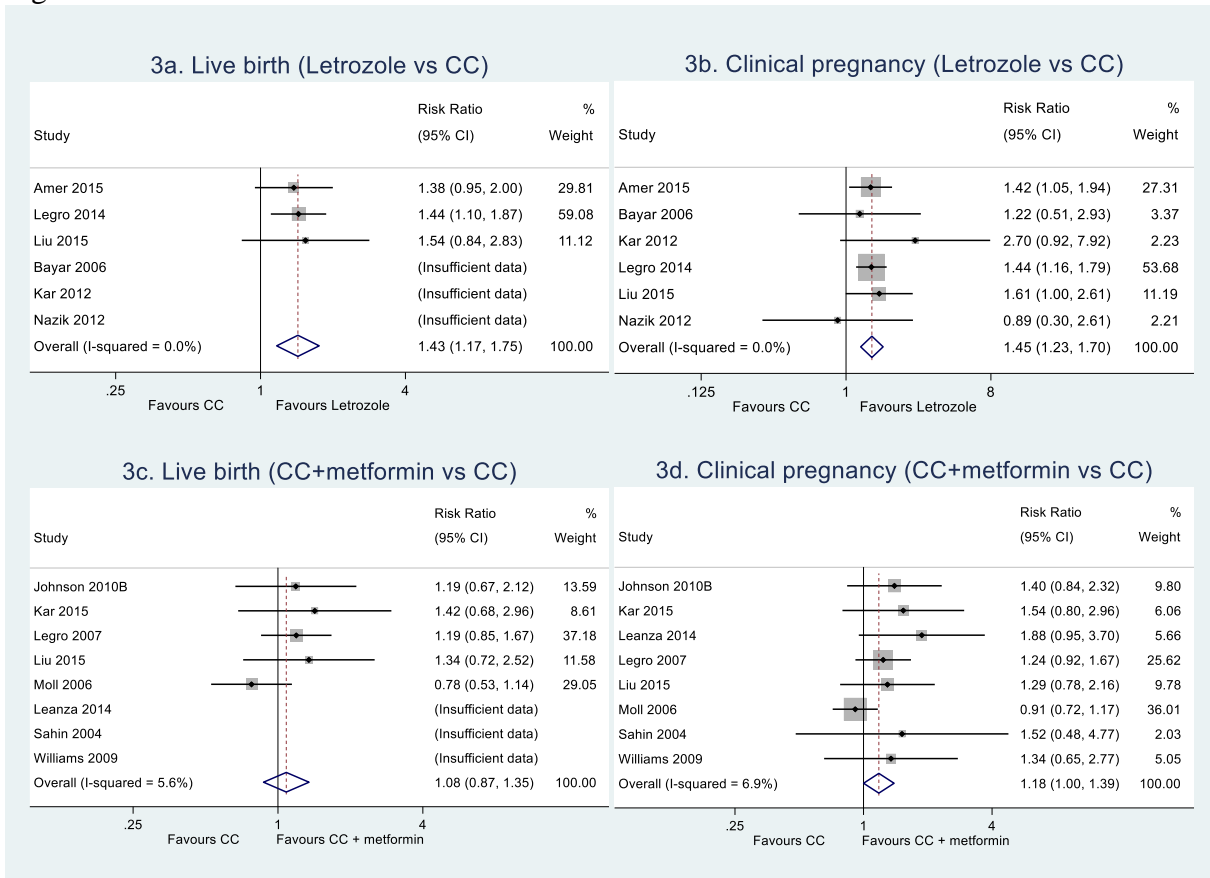
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912 Fig 2
 913

	Random sequence generation (selection bias)	Allocation concealment (selection bias)	Blinding of participants and personnel (performance bias)	Blinding of outcome assessment (detection bias)	Incomplete outcome data (attrition bias)	Selective reporting (reporting bias)	Other bias
Amer 2017	+	+	+	+	+	+	+
Bayar 2006	+	+	+	+	+	?	+
Homburg 2012	+	+	-	+	+	+	+
Johnson 2010A	+	+	+	+	+	+	+
Johnson 2010B	+	+	+	+	+	+	+
Kar 2012	+	-	-	+	+	?	+
Kar 2015	+	-	-	+	-	+	+
Leanza 2014	+	+	+	+	+	?	+
Legro 2007	+	+	+	+	+	+	+
Legro 2014	+	+	+	+	+	+	+
Liu 2017	+	-	-	+	+	+	+
Lord 2006	+	+	+	+	+	+	-
Moll 2006	+	+	+	+	+	+	+
Morin-Papunen 2012	+	+	+	+	+	+	+
Nazik 2012	+	-	-	+	+	?	+
Palomba 2005	+	+	+	+	+	+	+
Sahin 2004	+	+	+	+	+	?	+
Vegetti 1999	+	+	-	+	+	?	+
Williams 2009	+	+	+	+	+	?	+
Wu 2017	+	+	+	+	+	+	+

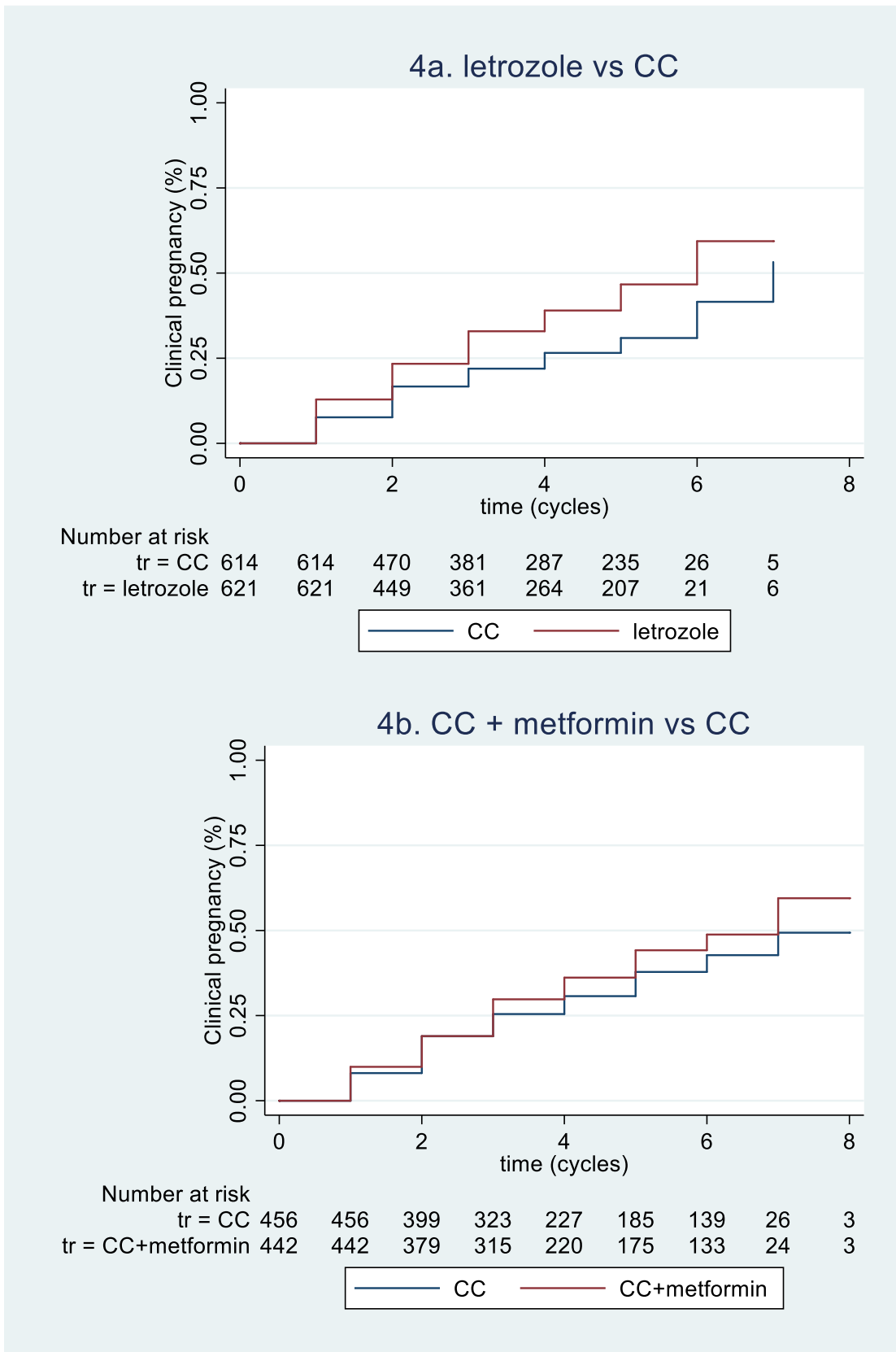
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917 Fig 3



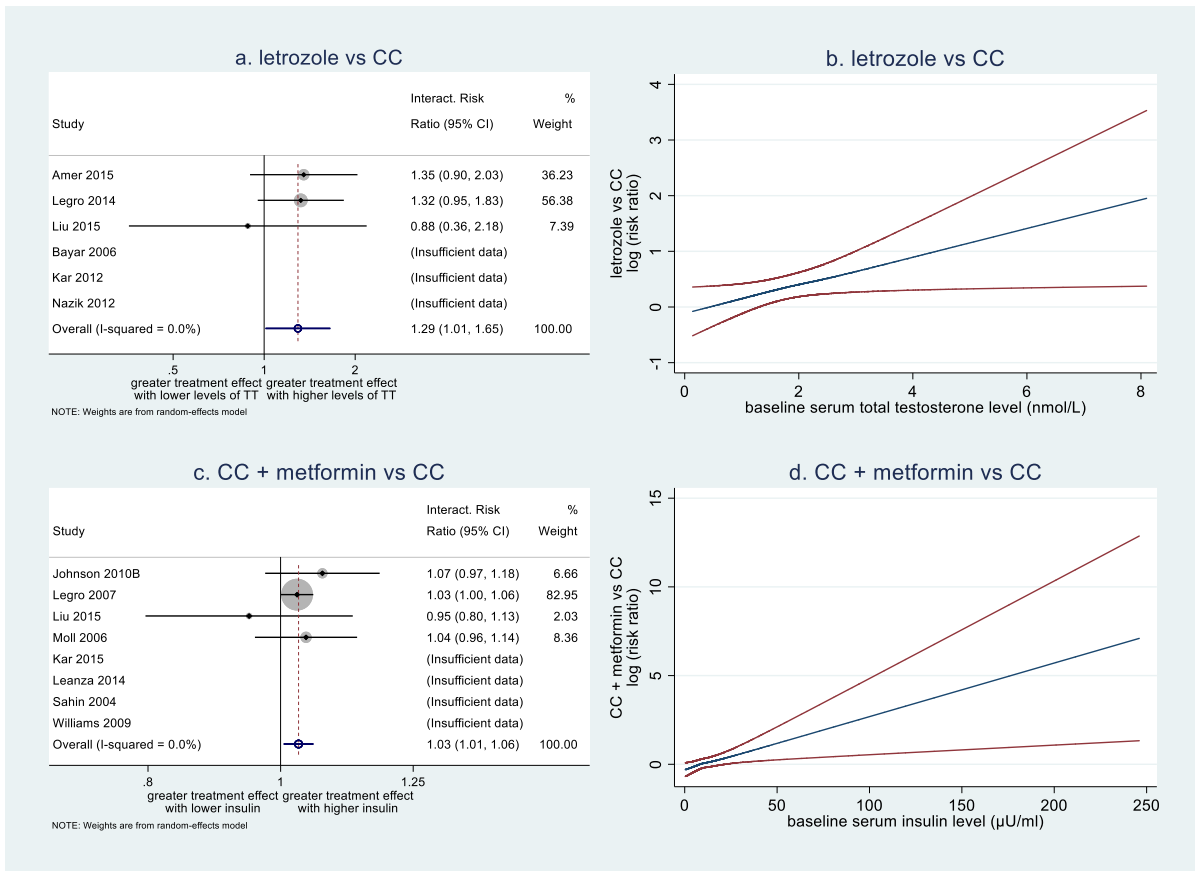
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920 Figure 4
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925 Figure 5
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